# Mechanisms for Generating Power-Law Size Distributions, Part 1

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Principles of Complex Systems, Vols. 1, 2, & 3D CSYS/MATH 6701, 6713, & a pretend number, 2024–2025

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The First Return Problem

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## Outline

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Heavy-tailed distributions are characters.

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Heavy-tailed distributions are **characters**.



Some of these distributions have power-law tails.

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measurers).

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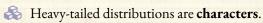
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Some of these distributions have power-law tails.

Measured exponents ( $\gamma$ 's and  $\alpha$ 's) vary across systems (and measurers).

What's their origin story?

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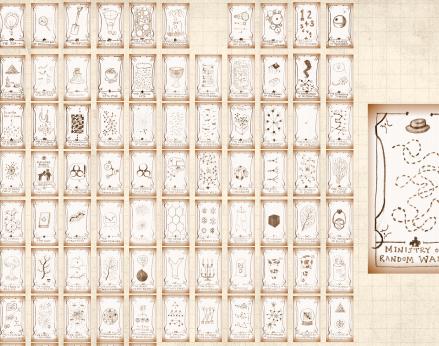
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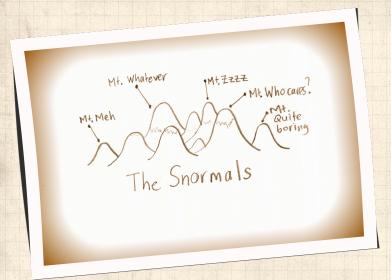
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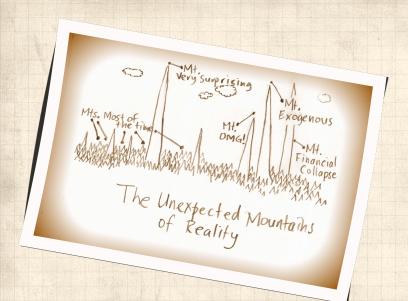
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A powerful story in the rise of complexity:

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A powerful story in the rise of complexity:



structure arises out of randomness.

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## A powerful story in the rise of complexity:



structure arises out of randomness.



& Exhibit A: Random walks.

#### The essential random walk:



One spatial dimension.

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& Exhibit A: Random walks.

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One spatial dimension.

Time and space are discrete

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One spatial dimension.

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Random walker (e.g., a zombie texter  $\Box$ ) starts at origin x = 0.

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## A powerful story in the rise of complexity:

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& Exhibit A: Random walks.

#### The essential random walk:

One spatial dimension.

Time and space are discrete

Random walker (e.g., a zombie texter  $\Box$ ) starts at origin x = 0.

 $\clubsuit$  Step at time t is  $\epsilon_t$ :

$$\epsilon_t = \left\{ \begin{array}{ll} +1 & \text{with probability } 1/2 \\ -1 & \text{with probability } 1/2 \end{array} \right.$$

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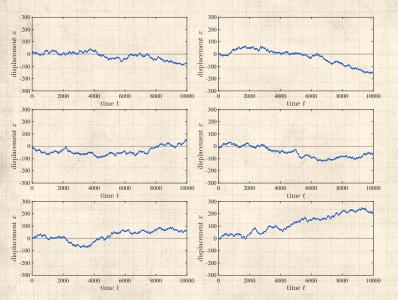
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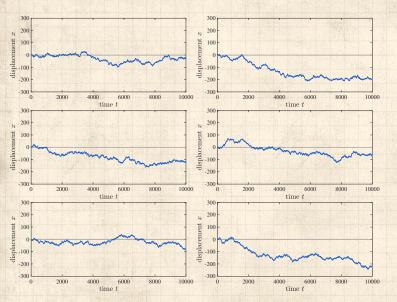
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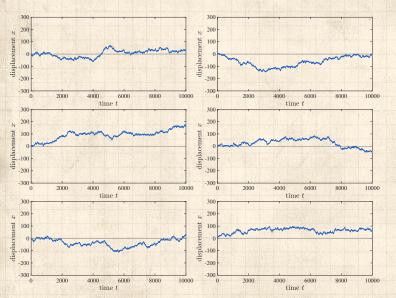
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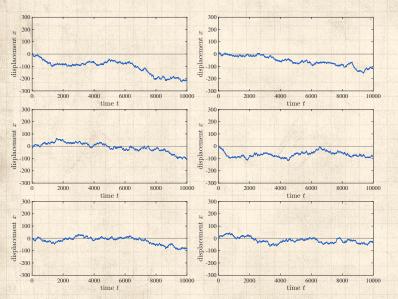
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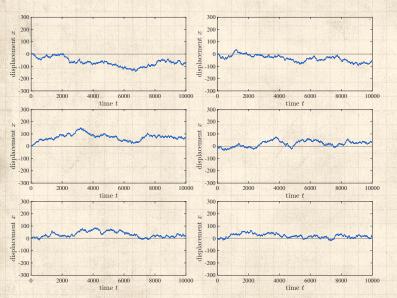
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## Displacement after t steps:

$$x_t = \sum_{i=1}^t \epsilon_i$$

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Displacement after t steps:

$$x_t = \sum_{i=1}^t \epsilon_i$$

Expected displacement:

$$\langle x_t \rangle = \left\langle \sum_{i=1}^t \epsilon_i \right\rangle$$

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Displacement after t steps:

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Expected displacement:

$$\langle x_t \rangle = \left\langle \sum_{i=1}^t \epsilon_i \right\rangle \\ = \sum_{i=1}^t \left\langle \epsilon_i \right\rangle$$

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At any time step, we 'expect' our zombie texter to be back at their starting place.

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- At any time step, we 'expect' our zombie texter to be back at their starting place.
- 💫 Obviously fails for odd number of steps...

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## Displacement after t steps:

$$x_t = \sum_{i=1}^t \epsilon_i$$

## Expected displacement:

$$\langle x_t \rangle = \left\langle \sum_{i=1}^t \epsilon_i \right\rangle \\ = \sum_{i=1}^t \left\langle \epsilon_i \right\rangle \\ = 0$$

- At any time step, we 'expect' our zombie texter to be back at their starting place.
- Obviously fails for odd number of steps...
- But as time goes on, the chance of our texting undead friend lurching back to x=0 must diminish, right?

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$$\mathrm{Var}(x_t) = \mathrm{Var}\left(\sum_{i=1}^t \epsilon_i\right)$$

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$$\begin{aligned} & \operatorname{Var}(x_t) = \operatorname{Var}\left(\sum_{i=1}^t \epsilon_i\right) \\ & = \sum_{i=1}^t \operatorname{Var}\left(\epsilon_i\right) \end{aligned}$$

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$$\begin{aligned} & \operatorname{Var}(x_t) = \operatorname{Var}\left(\sum_{i=1}^t \epsilon_i\right) \\ & = \sum_{i=1}^t \operatorname{Var}\left(\epsilon_i\right) = \sum_{i=1}^t 1 \end{aligned}$$

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$$\begin{aligned} & \operatorname{Var}(x_t) = \operatorname{Var}\left(\sum_{i=1}^t \epsilon_i\right) \\ & = \sum_{i=1}^t \operatorname{Var}\left(\epsilon_i\right) = \sum_{i=1}^t 1 = t \end{aligned}$$

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$$\begin{aligned} &\operatorname{Var}(x_t) = \operatorname{Var}\left(\sum_{i=1}^t \epsilon_i\right) \\ &= \sum_{i=1}^t \operatorname{Var}\left(\epsilon_i\right) = \sum_{i=1}^t 1 = t \end{aligned}$$

So typical displacement from the origin scales as:

$$\sigma = t^{1/2}$$

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\* Sum rule = a good reason for using the variance to measure spread; only works for independent distributions.

So typical displacement from the origin scales as:

$$\sigma = t^{1/2}$$

A non-trivial scaling law arises out of additive aggregation or accumulation.

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## Great moments in Televised Random Walks 2:



Plinko! Tfrom the Price is Right.

Plinko failure .



Also known as the bean machine , the quincunx (simulation) , and the Galton box.

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Counting random walks:

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## Counting random walks:



Each specific random walk of length t appears with a chance  $1/2^{t}$ .

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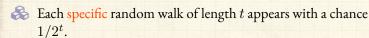
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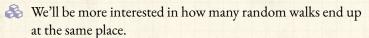
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## Counting random walks:





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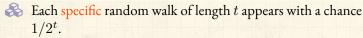
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#### Counting random walks:



We'll be more interested in how many random walks end up at the same place.

 $\Re$  Define N(i, j, t) as # distinct walks that start at x = i and end at x = j after t time steps.

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#### Counting random walks:

- Each specific random walk of length t appears with a chance  $1/2^t$ .
- We'll be more interested in how many random walks end up at the same place.
- $\Re$  Define N(i, j, t) as # distinct walks that start at x = i and end at x = j after t time steps.
- $\begin{cases} \& \& \end{cases}$  Random walk must displace by +(j-i) after t steps.

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- $\ensuremath{\mathfrak{S}}$  Random walk must displace by +(j-i) after t steps.
- insert assignment question

$$N(i,j,t) = \begin{pmatrix} t \\ (t+j-i)/2 \end{pmatrix}$$

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Take time t = 2n to help ourselves.

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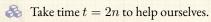
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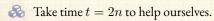
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$$x_{2n}$$
 is even so set  $x_{2n} = 2k$ .

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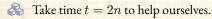
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$$x_{2n} \in \{0, \pm 2, \pm 4, \dots, \pm 2n\}$$

$$x_{2n}$$
 is even so set  $x_{2n} = 2k$ .

& Using our expression N(i,j,t) with i=0,j=2k, and t=2n, we have

$$\Pr(x_{2n} \equiv 2k) \propto \binom{2n}{n+k}$$

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For large *n*, the binomial deliciously approaches the Normal Distribution of Snoredom:

$$\Pr(x_t \equiv x) \simeq \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}}.$$

Insert assignment question

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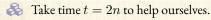
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The whole is different from the parts.

#nutritious

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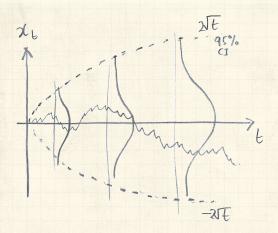
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References



See also: Stable Distributions

# Universality 🗗 is also not left-handed:



This is Diffusion 
 ∴: the most essential kind of spreading (more later).

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View as Random Additive Growth Mechanism.

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## Pascal's Triangle





Could have been the Pyramid of Pingala 10 or the Triangle of Khayyam, Jia Xian, Tartaglia, ...

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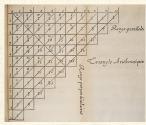
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Binomials tend towards the Normal.



## Pascal's Triangle





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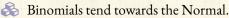
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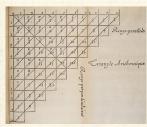
Counting encoded in algebraic forms (and much more).



¹Stigler's Law of Eponymy 

✓ showing excellent form again.

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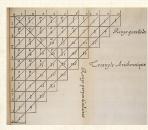


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<sup>&</sup>lt;sup>1</sup>Stigler's Law of Eponymy Showing excellent form again.

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Could have been the Pyramid of Pingala 1 or the Triangle of Khayyam, Jia Xian, Tartaglia, ...

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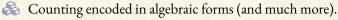
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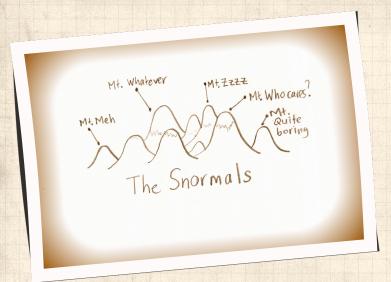


$$(h+t)^n = \sum_{k=0}^n {n \choose k} h^k t^{n-k} \text{ where } {n \choose k} = \frac{n!}{k!(n-k)!}$$



¹Stigler's Law of Eponymy 

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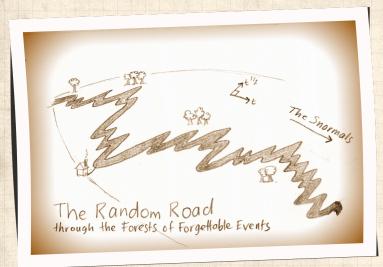
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## Random Walks

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 $\xi_{r,t}$  = the probability that by time step t, a random walk has crossed the origin r times.

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## Random Walks

The First Return

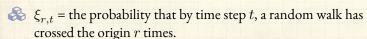
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Think of a coin flip game with ten thousand tosses.

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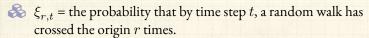
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Think of a coin flip game with ten thousand tosses.

If you are behind early on, what are the chances you will make a comeback? The PoCSverse Power-Law Mechanisms, Pt. 1 21 of 49

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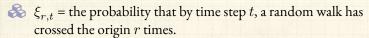
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Think of a coin flip game with ten thousand tosses.

If you are behind early on, what are the chances you will make a comeback?

The most likely number of lead changes is...

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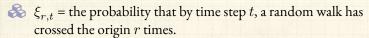
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Think of a coin flip game with ten thousand tosses.

If you are behind early on, what are the chances you will make a comeback?

The most likely number of lead changes is... 0.

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- $\xi_{r,t}$  = the probability that by time step t, a random walk has crossed the origin r times.
- Think of a coin flip game with ten thousand tosses.
- If you are behind early on, what are the chances you will make a comeback?
- The most likely number of lead changes is... 0.
- $\Re$  In fact:  $\xi_{0,t} > \xi_{1,t} > \xi_{2,t} > \cdots$

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- $\xi_{r,t}$  = the probability that by time step t, a random walk has crossed the origin r times.
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- Even crazier: The expected time between tied scores =  $\infty$

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- $\xi_{r,t}$  = the probability that by time step t, a random walk has crossed the origin r times.
- Think of a coin flip game with ten thousand tosses.
- If you are behind early on, what are the chances you will make a comeback?
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- & In fact:  $\xi_{0,t} > \xi_{1,t} > \xi_{2,t} > \cdots$
- & Even crazier:

The expected time between tied scores =  $\infty$ 

See Feller, Intro to Probability Theory, Volume I [5]

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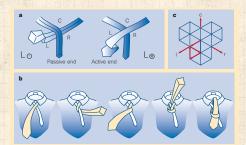
Fractional Brownian Motion



# Applied knot theory:



"Designing tie knots by random walks" Fink and Mao,
Nature, **398**, 31–32, 1999. [6]



**Rigure 1** All diagrams are drawn in the frame of reference of the mirror image of the actual ise. The two ways of beginning a knot,  $L_{\rm b}$  and  $L_{\rm b}$ . For knots beginning with  $L_{\rm c}$  the tie must begin inside-out,  $B_{\rm c}$ ,  $B_{\rm c}$  the formal denoted by the sequence  $L_{\rm c}$   $R_{\rm c}$   $L_{\rm c}$ ,  $C_{\rm c}$   $R_{\rm c}$  knot may be represented by a presistent random wells on a triangular lattice. The example shown is the four-in-hand, indicated by the wait 1118.

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# Applied knot theory:

Table 1 Aesthetic tie knots							
h	γ	γ/h	K(h, γ)	s	b	Name	Sequence
3	1	0.33	1	0	0		L <sub>o</sub> R <sub>⊗</sub> C <sub>o</sub> T
4	1	0.25	1	-1	1	Four-in-hand	$L_{\otimes}R_{\circ}L_{\otimes}C_{\circ}T$
5	2	0.40	2	-1	0	Pratt knot	$L_{\circ}C_{\otimes}R_{\circ}L_{\otimes}C_{\circ}T$
6	2	0.33	4	0	0	Half-Windsor	$L_{\otimes}R_{\circ}C_{\otimes}L_{\circ}R_{\otimes}C_{\circ}T$
7	2	0.29	6	-1	1		$L_{\circ}R_{\otimes}L_{\circ}C_{\otimes}R_{\circ}L_{\otimes}C_{\circ}T$
7	3	0.43	4	0	1		$L_{\circ}C_{\otimes}R_{\circ}C_{\otimes}L_{\circ}R_{\otimes}C_{\circ}T$
8	2	0.25	8	0	2		$L_{\otimes}R_{\circ}L_{\otimes}C_{\circ}R_{\otimes}L_{\circ}R_{\otimes}C_{\circ}T$
8	3	0.38	12	-1	0	Windsor	$L_{\otimes}C_{\circ}R_{\otimes}L_{\circ}C_{\otimes}R_{\circ}L_{\otimes}C_{\circ}T$
9	3	0.33	24	0	0		$L_{\circ}R_{\otimes}C_{\circ}L_{\otimes}R_{\circ}C_{\otimes}L_{\circ}R_{\otimes}C_{\circ}T$
9	4	0.44	8	-1	2		$L_{\circ}C_{\otimes}R_{\circ}C_{\otimes}L_{\circ}C_{\otimes}R_{\circ}L_{\otimes}C_{\circ}T$
Knote are characterized by half-winding number h, centre number w, centre fraction w/h, knote per class k/h w)							

symmetry s. balance b. name and sequence.



h = number of moves



 $\gamma = \text{number of center}$ moves



 $\Re K(h,\gamma) =$ 





 $b = \frac{1}{2} \sum_{i=2}^{h-1} |\omega_i + \omega_{i-1}|$ where  $\omega = \pm 1$  represents winding direction.

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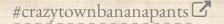
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## Random walks



The problem of first return:

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# Random Walks The First Return Problem

Random River Networks

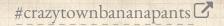
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## Random walks



#### The problem of first return:



Nhat is the probability that a random walker in one dimension returns to the origin for the first time after t steps? The PoCSverse Power-Law Mechanisms, Pt. 1 24 of 49 Random Walks

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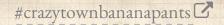
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#### The problem of first return:

Nill our zombie texter always return to the origin?

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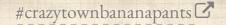
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#### The problem of first return:

 $\Leftrightarrow$  What is the probability that a random walker in one dimension returns to the origin for the first time after t steps?

Will our zombie texter always return to the origin?

What about higher dimensions?

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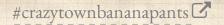
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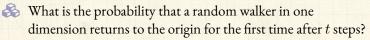
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#### The problem of first return:



Will our zombie texter always return to the origin?

What about higher dimensions?

Reasons for caring:

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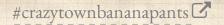
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#### The problem of first return:

Will our zombie texter always return to the origin?

What about higher dimensions?

#### Reasons for caring:

1. We will find a power-law size distribution with an interesting exponent.

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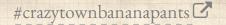
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#### The problem of first return:

What is the probability that a random walker in one dimension returns to the origin for the first time after t steps?

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What about higher dimensions?

### Reasons for caring:

- 1. We will find a power-law size distribution with an interesting exponent.
- 2. Some physical structures may result from random walks.

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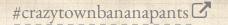
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#### The problem of first return:

 $\Leftrightarrow$  What is the probability that a random walker in one dimension returns to the origin for the first time after t steps?

Will our zombie texter always return to the origin?

What about higher dimensions?

### Reasons for caring:

- 1. We will find a power-law size distribution with an interesting exponent.
- 2. Some physical structures may result from random walks.
- 3. We'll start to see how different scalings relate to each other.

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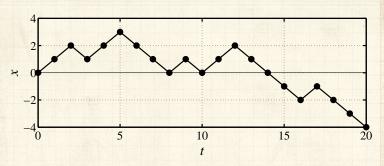
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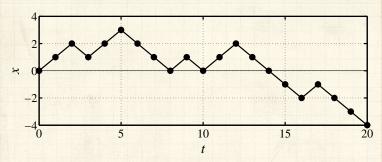
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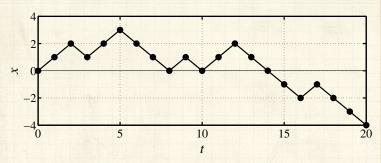
Death and Sports Fractional Brownian

Motion References



A return to origin can only happen when t = 2n.





A return to origin can only happen when t = 2n.

 $\clubsuit$  In example above, returns occur at t = 8, 10, and 14.

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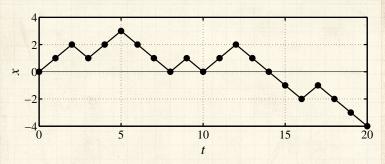
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A return to origin can only happen when t = 2n.

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 $\Leftrightarrow$  Call  $P_{\rm fr}(2n)$  the probability of first return at t=2n.

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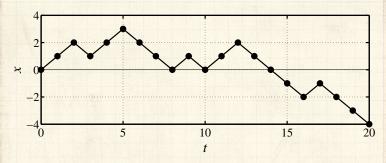
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Probability calculation 

≡ Counting problem (combinatorics/statistical mechanics).

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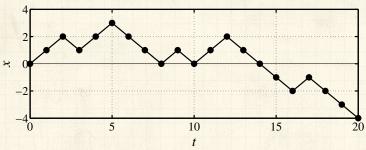
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 Idea: Transform first return problem into an easier return problem.

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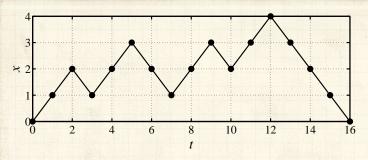
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Can assume zombie texter first lurches to x = 1.

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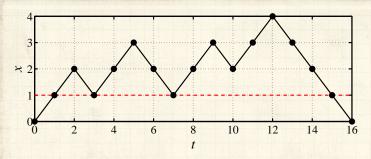
The First Return Problem

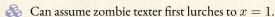
Random River Networks

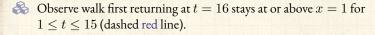
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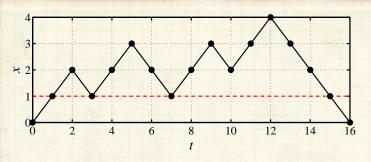
The First Return Problem

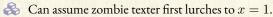
Random River Networks

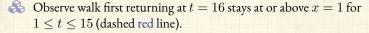
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Now want walks that can return many times to x = 1.

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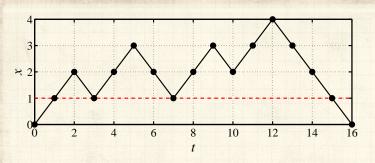
Problem

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- Can assume zombie texter first lurches to x = 1.
- Observe walk first returning at t=16 stays at or above x=1 for  $1 \le t \le 15$  (dashed red line).
- Now want walks that can return many times to x = 1.

Random Walks

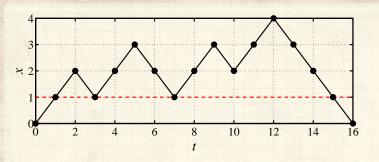
The First Return Problem

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- Observe walk first returning at t=16 stays at or above x=1 for  $1 \le t \le 15$  (dashed red line).
- Now want walks that can return many times to x = 1.
- $\begin{array}{l} \iff P_{\mathrm{fr}}(2n) = \\ 2 \cdot \frac{1}{2} Pr(x_t \geq 1, 1 \leq t \leq 2n-1, \text{ and } x_1 = x_{2n-1} = 1) \end{array}$
- $\mbox{\&}$  The  $\frac{1}{2}$  accounts for  $x_{2n}=2$  instead of 0.

Random Walks
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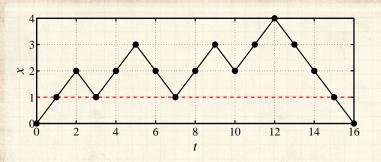
Problem

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- $\mbox{\&}$  The  $\frac{1}{2}$  accounts for  $x_{2n}=2$  instead of 0.
- $\clubsuit$  The 2 accounts for texters that first lurch to x = -1.

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Approach:

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#### Approach:



Move to counting numbers of walks.

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#### Approach:



Move to counting numbers of walks.



Return to probability at end.

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### Approach:



Move to counting numbers of walks.



Return to probability at end.



Again, N(i, j, t) is the # of possible walks between x = i and x = j taking t steps.

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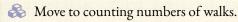
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### Approach:



Return to probability at end.

Again, N(i, j, t) is the # of possible walks between x = i and x = j taking t steps.

 $\ref{eq:consider}$  Consider all paths starting at x=1 and ending at x=1 after t=2n-2 steps.

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#### Approach:

- Move to counting numbers of walks.
- Return to probability at end.
- Again, N(i, j, t) is the # of possible walks between x = i and x = j taking t steps.
- $\ref{eq:consider}$  Consider all paths starting at x=1 and ending at x=1 after t=2n-2 steps.
- All Idea: If we can compute the number of walks that hit x=0 at least once, then we can subtract this from the total number to find the ones that maintain  $x \ge 1$ .

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### Approach:

- Move to counting numbers of walks.
- Return to probability at end.
- Again, N(i, j, t) is the # of possible walks between x = i and x = j taking t steps.
- $\ref{Solution}$  Consider all paths starting at x=1 and ending at x=1 after t=2n-2 steps.
- All Idea: If we can compute the number of walks that hit x=0 at least once, then we can subtract this from the total number to find the ones that maintain  $x \ge 1$ .
- $\ensuremath{\mathfrak{S}}$  Call walks that drop below x=1 excluded walks.

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### Approach:

- Move to counting numbers of walks.
- Return to probability at end.
- Again, N(i, j, t) is the # of possible walks between x = i and x = j taking t steps.
- $\Leftrightarrow$  Consider all paths starting at x=1 and ending at x=1 after t=2n-2 steps.
- All Idea: If we can compute the number of walks that hit x=0 at least once, then we can subtract this from the total number to find the ones that maintain  $x \ge 1$ .
- We'll use a method of images to identify these excluded walks.

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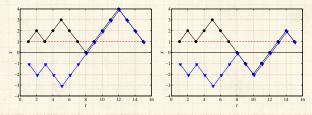
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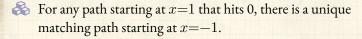
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# Key observation for excluded walks:



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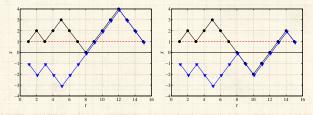
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### Key observation for excluded walks:

- For any path starting at x=1 that hits 0, there is a unique matching path starting at x=-1.
- Matching path first mirrors and then tracks after first reaching x=0.

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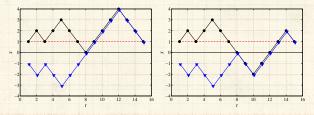
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## Key observation for excluded walks:

- For any path starting at x=1 that hits 0, there is a unique matching path starting at x=-1.
- Matching path first mirrors and then tracks after first reaching x=0.
- $\clubsuit$  # of t-step paths starting and ending at x=1 and hitting x=0 at least once

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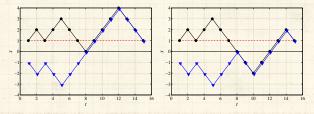
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## Key observation for excluded walks:

- For any path starting at x=1 that hits 0, there is a unique matching path starting at x=-1.
- Matching path first mirrors and then tracks after first reaching x=0.
- # of t-step paths starting and ending at x=1 and hitting x=0 at least once
  - = # of t-step paths starting at x=-1 and ending at x=+1

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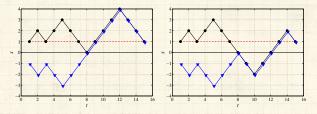
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Motion Brownian





## Key observation for excluded walks:

- For any path starting at x=1 that hits 0, there is a unique matching path starting at x=-1.
- Matching path first mirrors and then tracks after first reaching x=0.
- $\implies$  # of t-step paths starting and ending at x=1 and hitting x=0 at least once

= # of t-step paths starting at x=-1 and ending at x=+1 = N(-1, +1, t)

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 $\ref{eq:continuous}$  Call the number of paths that return after t=2n time steps after first moving to the positive side  $N_{\rm fr}^+(2n)$ .

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after first moving to the positive side  $N_{\rm fr}^+(2n)$ .



 $\mathfrak{F}$  For paths that first move to the negative side:  $N_{\mathrm{fr}}^{-}(2n)$ .

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after first moving to the positive side  $N_{\epsilon}^+(2n)$ .



 $\mathfrak{F}$  For paths that first move to the negative side:  $N_{\mathrm{fr}}^{-}(2n)$ .

 $\Re So N_{fr}^+(2n) = N(+1,+1,2n-2) - N(-1,+1,2n-2)$ 



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 Call the number of paths that return after t=2n time steps after first moving to the positive side  $N_{\epsilon}^+(2n)$ .



 $\begin{cases} \begin{cases} \begin{cases}$ 

 $\Re So N_{fr}^+(2n) = N(+1,+1,2n-2) - N(-1,+1,2n-2)$ 

Negative side:

$$N_{\text{fr}}^{-}(2n) = N(-1, -1, 2n - 2) - N(+1, -1, 2n - 2)$$

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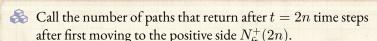
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 $\ensuremath{\mathfrak{S}}$  For paths that first move to the negative side:  $N_{
m fr}^-(2n)$ .

$$\Re$$
 So  $N_{\text{fr}}^+(2n) = N(+1, +1, 2n-2) - N(-1, +1, 2n-2)$ 

Negative side:  $N^{-}(2m) = N(-1)$ 

$$N_{\mathrm{fr}}^-(2n) = N(-1,-1,2n-2) - N(+1,-1,2n-2)$$

Symmetry:  $N_{\rm fr}^+(2n) = N_{\rm fr}^-(2n)$ 

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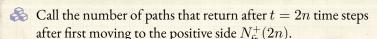
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 So  $N_{\text{fr}}^+(2n) = N(+1, +1, 2n-2) - N(-1, +1, 2n-2)$ 

Negative side:

$$N_{\mathrm{fr}}^{-}(2n) = N(-1, -1, 2n - 2) - N(+1, -1, 2n - 2)$$

Symmetry:  $N_{\rm fr}^{+}(2n) = N_{\rm fr}^{-}(2n)$ 

 $\ensuremath{\mathfrak{S}}$  Both  $N_{
m fr}(2n)$  and the one sided  $N_{
m fr}^+(2n)$  are of mathematical and physical interest.

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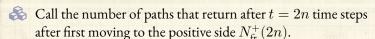
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 $\red{solution}$  For paths that first move to the negative side:  $N_{\mathrm{fr}}^-(2n)$ .

$$\Re$$
 So  $N_{\text{fr}}^+(2n) = N(+1, +1, 2n-2) - N(-1, +1, 2n-2)$ 

Negative side:

$$N_{\rm fr}^-(2n) = N(-1, -1, 2n - 2) - N(+1, -1, 2n - 2)$$

Symmetry:  $N_{\rm fr}^{+}(2n) = N_{\rm fr}^{-}(2n)$ 

 $\ \,$  Both  $N_{\rm fr}(2n)$  and the one sided  $N_{\rm fr}^+(2n)$  are of mathematical and physical interest.

A Overall:

$$\begin{split} N_{\mathrm{fr}}(2n) &= N_{\mathrm{fr}}^+(2n) + N_{\mathrm{fr}}^-(2n) = 2N_{\mathrm{fr}}^+(2n) \\ &= 2N(+1,+1,2n-2) - 2N(-1,+1,2n-2). \end{split}$$

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Insert assignment question 2:

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Insert assignment question 2:



$$N_{
m fr}(2n) \sim rac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}.$$

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Insert assignment question 2:



$$N_{
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Normalized number of paths gives probability.

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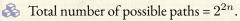


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Insert assignment question 2:



$$N_{\rm fr}(2n) \sim rac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}.$$

Normalized number of paths gives probability.

Total number of possible paths =  $2^{2n}$ .



$$P_{\rm fr}(2n) = \frac{1}{2^{2n}} N_{\rm fr}(2n)$$

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### Insert assignment question 2:



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Normalized number of paths gives probability.

Total number of possible paths =  $2^{2n}$ .

8

$$P_{\rm fr}(2n) = \frac{1}{2^{2n}} N_{\rm fr}(2n)$$
$$\simeq \frac{1}{2^{2n}} \frac{2^{2n-3/2}}{\sqrt{2\pi} n^{3/2}}$$

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### Insert assignment question 2:



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 $\Leftrightarrow$  We have  $P(t) \propto t^{-3/2}$ ,  $\gamma = 3/2$ .

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 $\Leftrightarrow$  We have  $P(t) \propto t^{-3/2}$ ,  $\gamma = 3/2$ .



Same scaling holds for continuous space/time walks.

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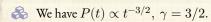
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Same scaling holds for continuous space/time walks.

P(t) is normalizable.

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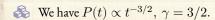
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Same scaling holds for continuous space/time walks.

 $\Re P(t)$  is normalizable.

Recurrence: Random walker always returns to origin

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But mean, variance, and all higher moments are infinite. #totalmadness The PoCSverse Power-Law Mechanisms, Pt. 1 31 of 49

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& Even though walker must return, expect a long wait...

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- One moral: Repeated gambling against an infinitely wealthy opponent must lead to ruin.

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Higher dimensions 2:

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# Higher dimensions 2:

Walker in d=2 dimensions must also return

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# Higher dimensions 2:

- Walker in d=2 dimensions must also return
- $\mbox{\&}$  Walker may not return in  $d \geq 3$  dimensions

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# Higher dimensions 2:

- Walker in d=2 dimensions must also return
- $\mbox{\&}$  Walker may not return in  $d \geq 3$  dimensions
- Associated human genius: George Pólya 🗹

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On finite spaces:

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### On finite spaces:



In any finite homogeneous space, a random walker will visit every site with equal probability

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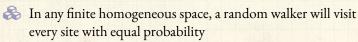
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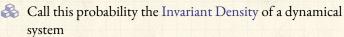
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### On finite spaces:





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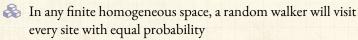
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### On finite spaces:



Call this probability the Invariant Density of a dynamical system

Non-trivial Invariant Densities arise in chaotic systems.

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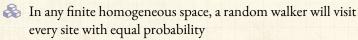
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On networks:

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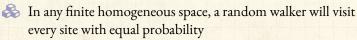
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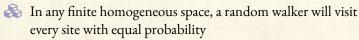
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References

#groovy



### On finite spaces:



Call this probability the Invariant Density of a dynamical system

Non-trivial Invariant Densities arise in chaotic systems.

On networks:

On networks, a random walker visits each node with frequency 

node degree #groovy

Equal probability still present: walkers traverse edges with equal frequency.

#totallygroovy

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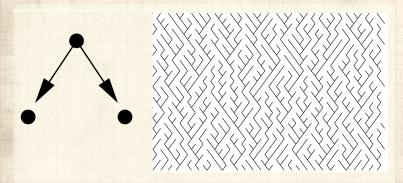
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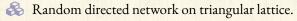
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# Scheidegger Networks [17,4]





Toy model of real networks.

'Flow' is southeast or southwest with equal probability.

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Creates basins with random walk boundaries.

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Creates basins with random walk boundaries.



Solution Observe that subtracting one random walk from another gives random walk with increments:

$$\epsilon_t = \left\{ \begin{array}{ll} +1 & \text{with probability } 1/4 \\ 0 & \text{with probability } 1/2 \\ -1 & \text{with probability } 1/4 \end{array} \right.$$

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Creates basins with random walk boundaries.



Note: The contracting one random walk from another gives random walk with increments:

$$\epsilon_t = \left\{ \begin{array}{ll} +1 & \text{with probability } 1/4 \\ 0 & \text{with probability } 1/2 \\ -1 & \text{with probability } 1/4 \end{array} \right.$$



Random walk with probabilistic pauses.

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Random walk with probabilistic pauses.



Basin termination = first return random walk problem.

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Random walk with probabilistic pauses.



Basin termination = first return random walk problem.

Basin length  $\ell$  distribution:  $P(\ell) \propto \ell^{-3/2}$ 

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Creates basins with random walk boundaries.



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Random walk with probabilistic pauses.



Basin termination = first return random walk problem.





 $\Leftrightarrow$  For real river networks, generalize to  $P(\ell) \propto \ell^{-\gamma}$ .

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 $\red$  For a basin of length  $\ell$ , width  $\propto \ell^{1/2}$ 

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 $\Leftrightarrow$  For a basin of length  $\ell$ , width  $\propto \ell^{1/2}$ 



 $\begin{cases} \& \end{cases} \begin{cases} \begin{cases} \text{Basin area } a \propto \ell \cdot \ell^{1/2} = \ell^{3/2} \end{cases}$ 

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 $\ref{heliable}$  For a basin of length  $\ell$ , width  $\propto \ell^{1/2}$ 



 $\red Basin area \ a \propto \ell \cdot \ell^{1/2} = \ell^{3/2}$ 



A Invert:  $\ell \propto a^{2/3}$ 

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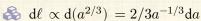
 $\ref{height}$  For a basin of length  $\ell$ , width  $\propto \ell^{1/2}$ 



 $\red{abs}$  Basin area  $a \propto \ell \cdot \ell^{1/2} = \ell^{3/2}$ 



A Invert:  $\ell \propto a^{2/3}$ 



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Solution For a basin of length  $\ell$ , width  $\propto \ell^{1/2}$ 



 $\clubsuit$  Basin area  $a \propto \ell \cdot \ell^{1/2} = \ell^{3/2}$ 



A Invert:  $\ell \propto a^{2/3}$ 



 $4 \otimes d\ell \propto d(a^{2/3}) = 2/3a^{-1/3}da$ 



 $\operatorname{\mathbf{Pr}}(\operatorname{basin} \operatorname{area} = a) da$ = **Pr**(basin length  $= \ell$ )d $\ell$  The PoCSverse Power-Law Mechanisms, Pt. 1 35 of 49

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So For a basin of length  $\ell$ , width  $\propto \ell^{1/2}$ 



 $\clubsuit$  Basin area  $a \propto \ell \cdot \ell^{1/2} = \ell^{3/2}$ 



A Invert:  $\ell \propto a^{2/3}$ 



 $d\ell \propto d(a^{2/3}) = 2/3a^{-1/3}da$ 



 $\operatorname{\mathbf{Pr}}(\operatorname{basin} \operatorname{area} = a) da$  $= \mathbf{Pr}(\text{basin length} = \ell) d\ell$  $\propto \ell^{-3/2} d\ell$ 

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Solution For a basin of length  $\ell$ , width  $\propto \ell^{1/2}$ 



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A Invert:  $\ell \propto a^{2/3}$ 



 $d\ell \propto d(a^{2/3}) = 2/3a^{-1/3}da$ 



 $\mathbf{R}$  **Pr**(basin area = a)da = **Pr**(basin length  $= \ell$ )d $\ell$  $\propto \ell^{-3/2} d\ell$  $\propto (a^{2/3})^{-3/2}a^{-1/3}da$ 

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So For a basin of length  $\ell$ , width  $\propto \ell^{1/2}$ 



 $\clubsuit$  Basin area  $a \propto \ell \cdot \ell^{1/2} = \ell^{3/2}$ 



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Both basin area and length obey power law distributions

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Both basin area and length obey power law distributions

Observed for real river networks

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Reportedly:  $1.3 < \tau < 1.5$  and  $1.5 < \gamma < 2$ 

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Hack's law [10]:

 $\ell \propto a^h$ 

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Models exist with interesting values of h.

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Models exist with interesting values of h.

A Plan: Redo calc with  $\gamma$ ,  $\tau$ , and h.

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 $\ell \propto a^h, \ P(a) \propto a^{-\tau}, \ {\rm and} \ P(\ell) \propto \ell^{-\gamma}$ 

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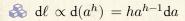
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$$\ell \propto a^h, \, P(a) \propto a^{-\tau}, \, {\rm and} \, P(\ell) \propto \ell^{-\gamma}$$

 $\Longrightarrow$  Find  $\tau$  in terms of  $\gamma$  and h.

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$$\ell \propto a^h, \, P(a) \propto a^{-\tau}, \, {\rm and} \, P(\ell) \propto \ell^{-\gamma}$$

 $\clubsuit$  Find  $\tau$  in terms of  $\gamma$  and h.

 $\mathbf{Pr}(\text{basin area} = a) da$   $= \mathbf{Pr}(\text{basin length} = \ell) d\ell$ 

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$$\ell \propto a^h, \, P(a) \propto a^{-\tau}, \, {\rm and} \, P(\ell) \propto \ell^{-\gamma}$$

 $\mbox{\&}$  Find au in terms of  $\gamma$  and h.

 $\begin{aligned} & \mathbf{Pr}(\text{basin area} = a) da \\ & = \mathbf{Pr}(\text{basin length} = \ell) d\ell \\ & \propto \ell^{-\gamma} d\ell \end{aligned}$ 

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$$\tau = 1 + h(\gamma - 1)$$

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$$\tau = 1 + h(\gamma - 1)$$

Excellent example of the Scaling Relations found between exponents describing power laws for many systems.

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With more detailed description of network structure,  $\tau = 1 + h(\gamma - 1)$  simplifies to: [3]

$$\tau = 2 - h$$

and

$$\gamma = 1/h$$

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With more detailed description of network structure,  $\tau = 1 + h(\gamma - 1)$  simplifies to: [3]

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 $\mathfrak{S}$  Only one exponent is independent (take h).

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Simplifies system description.

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Simplifies system description.



Expect Scaling Relations where power laws are found.

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With more detailed description of network structure,  $\tau=1+h(\gamma-1)$  simplifies to: <sup>[3]</sup>

$$\tau = 2 - h$$

and

$$\gamma = 1/h$$

- Only one exponent is independent (take h).
- Simplifies system description.
- Expect Scaling Relations where power laws are found.
- Need only characterize Universality class with independent exponents.

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### Death ...

#### Failure:



A very simple model of failure/death



 $x_t$  = entity's 'health' at time t



 $\Leftrightarrow$  Start with  $x_0 > 0$ .



"Explaining mortality rate plateaus"

Weitz and Fraser,

Proc. Natl. Acad. Sci., 98, 15383-15386, 2001. [18]

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### ... and the NBA:

### Basketball and other sports [2]:

Three arcsine laws  $\square$  (Lévy [12]) for continuous-time random walk lasting time T:

$$\frac{1}{\pi} \frac{1}{\sqrt{t(T-t)}}.$$

The arcsine distribution 2 applies for:

- (1) fraction of time positive,
- (2) the last time the walk changes sign, and (3) the time the maximum is achieved.
- $\ensuremath{\mathfrak{S}}$  Well approximated by basketball score lines  $^{[8,\,2]}$ .
- Australian Rules Football has some differences [11].

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& Can generalize to Fractional Random Walks [15, 16, 14]

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🚓 Fractional Brownian Motion 🗷, Lévy flights 🗹

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See Montroll and Shlesinger for example: [14] "On 1/f noise and other distributions with long tails." Proc. Natl. Acad. Sci., 1982.



 $\triangle$  In 1-d, standard deviation  $\sigma$  scales as

 $\sigma \sim t^{\alpha}$ 

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A Fractional Brownian Motion C, Lévy flights C



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 $\sigma \sim t^{\alpha}$ 

 $\alpha = 1/2$  — diffusive

 $\alpha > 1/2$  — superdiffusive

 $\alpha < 1/2$  — subdiffusive

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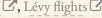




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Extensive memory of path now matters...

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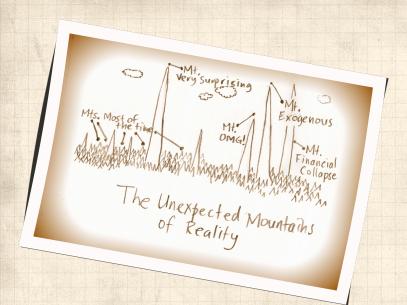
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- A First big studies of movement and interactions of people.
- & Brockmann *et al.* [1] "Where's George" study.
- Beyond Lévy: Superdiffusive in space but with long waiting times.
- $\ensuremath{\mathfrak{S}}$  Tracking movement via cell phones  $^{[9]}$  and Twitter  $^{[7]}$ .





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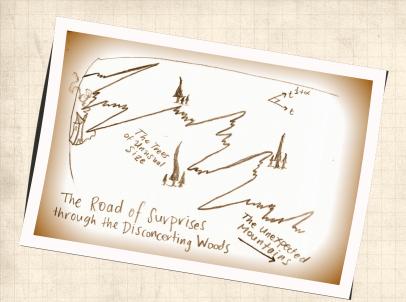
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