Mechanisms for Generating Power-Law Size Distributions, Part 1

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Principles of Complex Systems, Vols. 1, 2, & 3D CSYS/MATH 300, 303, & 394, 2022–2023 @pocsvox

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> The First Return Problem

The PoCSverse

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



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Outline

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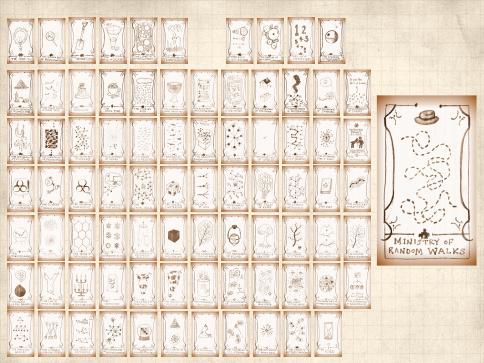
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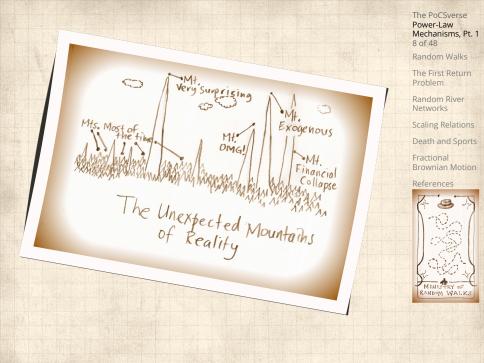
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A powerful story in the rise of complexity:

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A powerful story in the rise of complexity: structure arises out of randomness.



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A powerful story in the rise of complexity:
♣ structure arises out of randomness.
♣ Exhibit A: Random walks.

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A powerful story in the rise of complexity:
Structure arises out of randomness.
Structure A: Random walks. ☑

The essential random walk:

- 🚳 One spatial dimension.
- 🗞 Time and space are discrete

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The essential random walk:

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Random walker (e.g., a zombie texter \mathbb{C}) starts at origin x = 0.

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A powerful story in the rise of complexity:
Structure arises out of randomness.
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The essential random walk:

- 🚳 One spatial dimension.
- 🚳 Time and space are discrete
- Random walker (e.g., a zombie texter \square) starts at origin x = 0.
- \clubsuit Step at time t is ϵ_t :

 $\epsilon_t = \left\{ \begin{array}{l} +1 & \text{with probability 1/2} \\ -1 & \text{with probability 1/2} \end{array} \right.$

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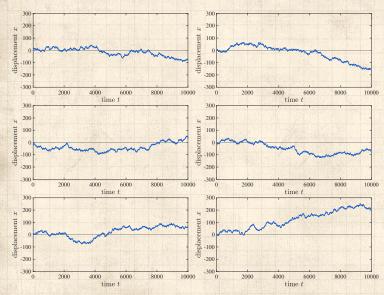
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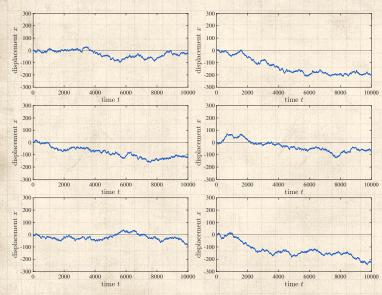
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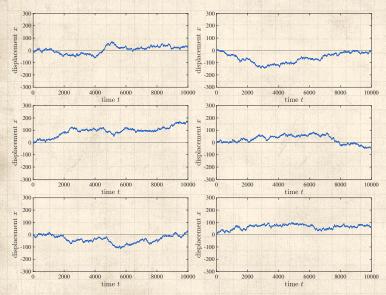
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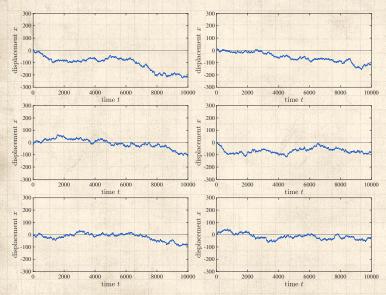
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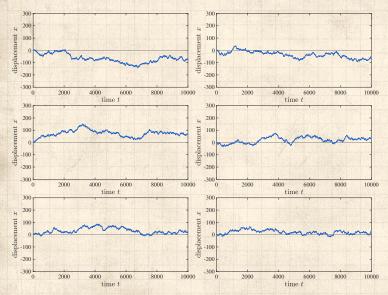
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Displacement after *t* steps:

$$x_t = \sum_{i=1}^t \epsilon_i$$

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Displacement after *t* steps:

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Expected displacement:

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At any time step, we 'expect' our zombie texter to be back at their starting place. The PoCSverse Power-Law Mechanisms, Pt. 1 11 of 48

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At any time step, we 'expect' our zombie texter to be back at their starting place.
 Obviously fails for odd number of steps...

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At any time step, we 'expect' our zombie texter to be back at their starting place.

- Obviously fails for odd number of steps...
- But as time goes on, the chance of our texting undead friend lurching back to x=0 must diminish, right?

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$$\mathsf{Var}(x_t) = \mathsf{Var}\left(\sum_{i=1}^t \epsilon_i\right)$$

* Sum rule = a good reason for using the variance to measure spread; only works for independent distributions.

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$$= \sum_{i=1}^t \operatorname{Var}\left(\epsilon_i\right)$$

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$$\mathsf{Var}(x_t) = \mathsf{Var}\left(\sum_{i=1}^t \epsilon_i\right)$$

$$=\sum_{i=1}^{\iota}\operatorname{Var}\left(\epsilon_{i}\right)=\sum_{i=1}^{\iota}1$$

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So typical displacement from the origin scales as:

$$\sigma = t^{1/2}$$

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A non-trivial scaling law arises out of additive aggregation or accumulation. The PoCSverse Power-Law Mechanisms, Pt. 1 12 of 48

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Great moments in Televised Random Walks:

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References



http://www.youtube.com/watch?v=05gqx6eSyO0?rel=0

Also known as the bean machine , the quincunx (simulation) , and the Galton box.

Random walk basics:

Counting random walks:

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Random walk basics:

Counting random walks:



Each specific random walk of length t appears with a chance $1/2^t$.

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Counting random walks:

- Each specific random walk of length t appears with a chance $1/2^t$.
- We'll be more interested in how many random walks end up at the same place.

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Counting random walks:

- Each specific random walk of length t appears with a chance $1/2^t$.
- We'll be more interested in how many random walks end up at the same place.
- Solution Define N(i, j, t) as # distinct walks that start at x = i and end at x = j after t time steps.

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- Random walk must displace by +(j-i) after t steps.

🚳 Insert question from assignment 5 🗹

$$N(i,j,t) = \binom{t}{(t+j-i)/2}$$

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 $rac{1}{8}$ Take time t = 2n to help ourselves.

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Solution Take time t = 2n to help ourselves. $x_{2n} \in \{0, \pm 2, \pm 4, \dots, \pm 2n\}$ The PoCSverse Power-Law Mechanisms, Pt. 1 15 of 48

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So Take time t = 2n to help ourselves. So $x_{2n} \in \{0, \pm 2, \pm 4, \dots, \pm 2n\}$ So x_{2n} is even so set $x_{2n} = 2k$. The PoCSverse Power-Law Mechanisms, Pt. 1 15 of 48

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How does $P(x_t)$ behave for large t? rightarrow Take time t = 2n to help ourselves.

- $\textcircled{S} x_{2n} \in \{0,\pm 2,\pm 4,\ldots,\pm 2n\}$
- x_{2n} is even so set $x_{2n} = 2k$.
- Solution Using our expression N(i, j, t) with i = 0, j = 2k, and t = 2n, we have

$$\Pr(x_{2n} \equiv 2k) \propto \binom{2n}{n+k}$$

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- rightarrow Take time t = 2n to help ourselves.
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For large n, the binomial deliciously approaches the Normal Distribution of Snoredom:

$$\label{eq:prod} \mathbf{Pr}(x_t \equiv x) \simeq \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}}.$$

Insert question from assignment 5 🖸

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For large n, the binomial deliciously approaches the Normal Distribution of Snoredom:

$$\label{eq:prime} \mathbf{Pr}(x_t \equiv x) \simeq \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}}.$$

Insert question from assignment 5 🗹 🗞 The whole is different from the parts.

#nutritious

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- $rac{1}{3}$ Take time t = 2n to help ourselves.
- $x_{2n} \in \{0, \pm 2, \pm 4, \dots, \pm 2n\}$
- x_{2n} is even so set $x_{2n} = 2k$.
- \bigotimes Using our expression N(i, j, t) with i = 0, j = 2k, and t = 2n, we have

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Insert question from assignment 5 🕑 The whole is different from the parts. #nutritious 😤 See also: Stable Distributions 🗹

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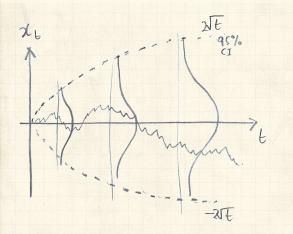
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Universality C is also not left-handed:



This is Diffusion C: the most essential kind of spreading (more later).

🚳 View as Random Additive Growth Mechanism.

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Pascal's Triangle





Could have been the Pvramid of Pingala C¹ or the Triangle of Khayyam, lia Xian, Tartaglia, ...

Binomials tend towards the Normal.

¹Stigler's Law of Eponymy C showing excellent form again.

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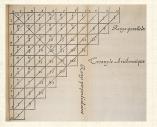
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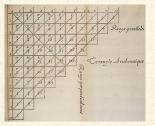


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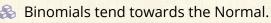
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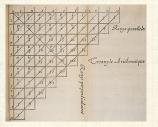
Counting encoded in algebraic forms (and much 2 more).

$$\bigotimes (h+t)^n = \sum_{k=0}^n {n \choose k} h^k t^{n-k}$$
 where ${n \choose k} = \frac{n!}{k!(n-k)!}$

ANDOM WALKS

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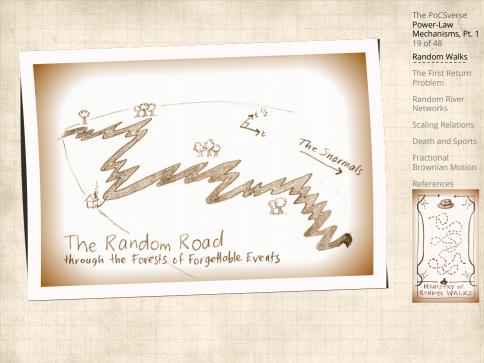
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 $\textcircled{h} (h+t)^3 = hhh + hht + hth + thh + htt + tht + tth + ttt$

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$\xi_{r,t}$ = the probability that by time step *t*, a random walk has crossed the origin *r* times.

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- $\underset{r,t}{\bigotimes} \xi_{r,t}$ = the probability that by time step t, a random walk has crossed the origin r times.
- A Think of a coin flip game with ten thousand tosses.

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- 🚓 The most likely number of lead changes is...

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- $\underset{r,t}{\Leftrightarrow} \xi_{r,t}$ = the probability that by time step t, a random walk has crossed the origin r times.
- A Think of a coin flip game with ten thousand tosses.
- If you are behind early on, what are the chances you will make a comeback?
- The most likely number of lead changes is... 0.
- & In fact: $\xi_{0,t} > \xi_{1,t} > \xi_{2,t} > \cdots$

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- $\underset{r,t}{\Leftrightarrow} \xi_{r,t}$ = the probability that by time step t, a random walk has crossed the origin r times.
- A Think of a coin flip game with ten thousand tosses.
- If you are behind early on, what are the chances you will make a comeback?
- A The most likely number of lead changes is... 0.

$$\diamondsuit$$
 In fact: $\xi_{0,t} > \xi_{1,t} > \xi_{2,t} > \cdot$

\lambda Even crazier:

The expected time between tied scores = ∞

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See Feller, Intro to Probability Theory, Volume I^[5]

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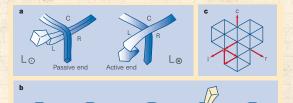
Fractional Brownian Motion

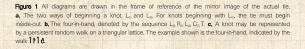


Applied knot theory:



"Designing tie knots by random walks" Fink and Mao, Nature, **398**, 31–32, 1999.^[6]





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Applied knot theory:

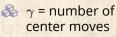
Table 1 Aesthetic tie knots

- 2							
h	γ	γ/h	$K(h, \gamma)$	S	b	Name	Sequence
3	1	0.33	1	0	0		L _☉ R _⊗ C _☉ T
4	1	0.25	1	- 1	1	Four-in-hand	L _⊗ R _☉ L _⊗ C _☉ T
5	2	0.40	2	- 1	0	Pratt knot	$L_{\odot}C_{\otimes}R_{\odot}L_{\otimes}C_{\odot}T$
6	2	0.33	4	0	0	Half-Windsor	$L_{\otimes}R_{\odot}C_{\otimes}L_{\odot}R_{\otimes}C_{\odot}T$
7	2	0.29	6	- 1	1		$L_{\odot}R_{\otimes}L_{\odot}C_{\otimes}R_{\odot}L_{\otimes}C_{\odot}T$
7	3	0.43	4	0	1		$L_{\odot}C_{\otimes}R_{\odot}C_{\otimes}L_{\odot}R_{\otimes}C_{\odot}T$
8	2	0.25	8	0	2		$L_{\otimes}R_{\odot}L_{\otimes}C_{\odot}R_{\otimes}L_{\odot}R_{\otimes}C_{\odot}T$
8	3	0.38	12	- 1	0	Windsor	$L_{\otimes}C_{\odot}R_{\otimes}L_{\odot}C_{\otimes}R_{\odot}L_{\otimes}C_{\odot}T$
9	3	0.33	24	0	0		$L_{\circ}R_{\otimes}C_{\circ}L_{\otimes}R_{\circ}C_{\otimes}L_{\circ}R_{\otimes}C_{\circ}T$
9	4	0.44	8	- 1	2		$L_{\odot}C_{\otimes}R_{\odot}C_{\otimes}L_{\odot}C_{\otimes}R_{\odot}L_{\otimes}C_{\odot}T$

Knots are characterized by half-winding number h, centre number γ , centre fraction γ/h , knots per class $K(h, \gamma)$, symmetry s, balance b, name and sequence.

5

h = number of moves



 $K(h,\gamma) = 2^{\gamma-1} \binom{h-\gamma-2}{\gamma-1}$

 $s = \sum_{i=1}^{h} x_i \text{ where } x = -1$ for *L* and +1 for *R*.

$$b = \frac{1}{2} \sum_{i=2}^{h-1} |\omega_i + \omega_{i-1}|$$

where $\omega = \pm 1$
represents winding
direction.

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The problem of first return:

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The problem of first return:

What is the probability that a random walker in one dimension returns to the origin for the first time after t steps? The PoCSverse Power-Law Mechanisms, Pt. 1 23 of 48

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The problem of first return:

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Will our zombie texter always return to the origin?

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The problem of first return:

- What is the probability that a random walker in one dimension returns to the origin for the first time after t steps?
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Reasons for caring:

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Random walks #crazytownbananapants

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Random walks #crazytownbananapants

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Reasons for caring:

- 1. We will find a power-law size distribution with an interesting exponent.
- 2. Some physical structures may result from random walks.
- 3. We'll start to see how different scalings relate to each other.

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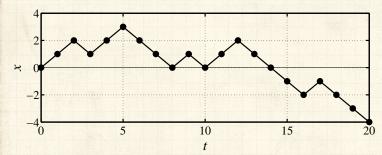
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For random walks in 1-d:



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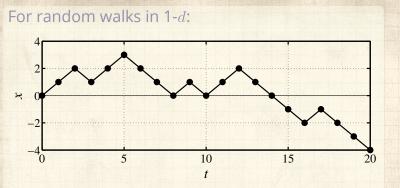
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\Re A return to origin can only happen when t = 2n.

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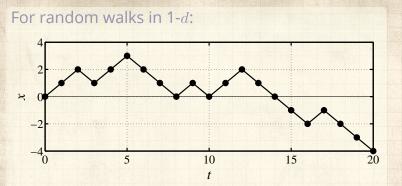
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A return to origin can only happen when t = 2n. In example above, returns occur at t = 8, 10, and 14. The PoCSverse Power-Law Mechanisms, Pt. 1 24 of 48

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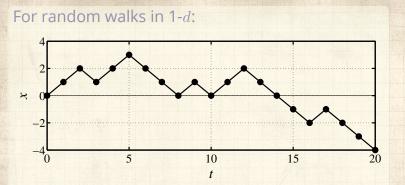
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A return to origin can only happen when t = 2n. In example above, returns occur at t = 8, 10, and 14.

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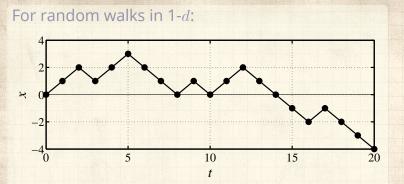
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- A return to origin can only happen when t = 2n. In example above, returns occur at t = 8, 10, and 14.
- \mathfrak{S} Call $P_{\mathsf{fr}(2n)}$ the probability of first return at t = 2n.
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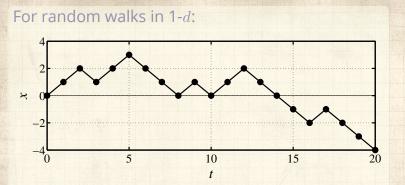
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- \mathfrak{R} Call $P_{\mathsf{fr}(2n)}$ the probability of first return at t = 2n.
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- Idea: Transform first return problem into an easier return problem.

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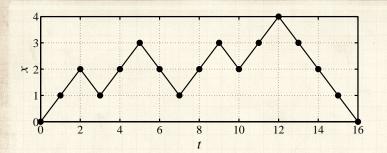
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 \bigotimes Can assume zombie texter first lurches to x = 1.

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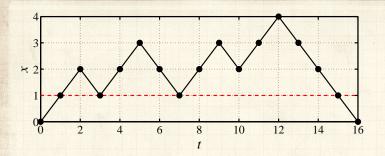
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Solution Observe walk first returning at t = 16 stays at or above x = 1 for $1 \le t \le 15$ (dashed red line).

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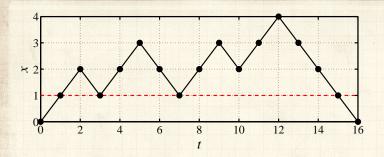
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Solution Can assume zombie texter first lurches to x = 1.

- Solution Observe walk first returning at t = 16 stays at or above x = 1 for $1 \le t \le 15$ (dashed red line).
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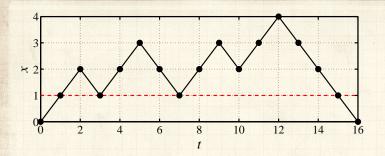
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 $\begin{array}{l} \bigotimes \ \ P_{\rm fr}(2n) = \\ 2 \cdot \frac{1}{2} Pr(x_t \geq 1, 1 \leq t \leq 2n-1, \text{ and } x_1 = x_{2n-1} = 1) \end{array} \\ \end{array}$

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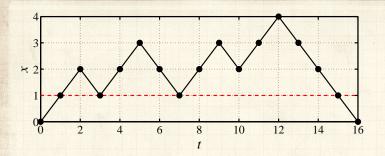
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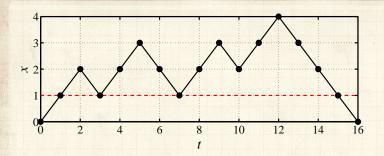
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 $rac{1}{2}$ The $rac{1}{2}$ accounts for $x_{2n} = 2$ instead of 0.

 \Im The 2 accounts for texters that first lurch to x = -1.

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Approach:

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Approach:



Move to counting numbers of walks.

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Approach:



Move to counting numbers of walks.

🚳 Return to probability at end.

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Approach:

- Move to counting numbers of walks.
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- Again, N(i, j, t) is the # of possible walks between x = i and x = j taking t steps.

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Approach:

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- Again, N(i, j, t) is the # of possible walks between x = i and x = j taking t steps.
- Solution Consider all paths starting at x = 1 and ending at x = 1 after t = 2n 2 steps.

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- Solution If we can compute the number of walks that hit x = 0 at least once, then we can subtract this from the total number to find the ones that maintain $x \ge 1$.

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- Solution Call walks that drop below x = 1 excluded walks.

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- rightarrow Call walks that drop below x = 1 excluded walks.
- We'll use a method of images to identify these excluded walks.

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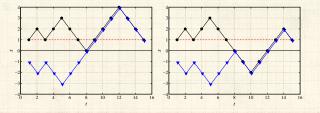
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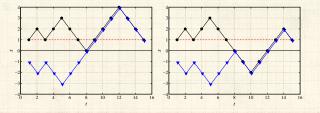
Fractional Brownian Motion

References



Key observation for excluded walks:

So For any path starting at x=1 that hits 0, there is a unique matching path starting at x=-1.



Key observation for excluded walks:

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- Matching path first mirrors and then tracks after first reaching x=0.

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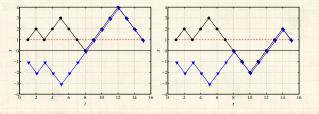
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Key observation for excluded walks:

- So For any path starting at x=1 that hits 0, there is a unique matching path starting at x=-1.
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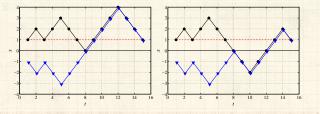
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of *t*-step paths starting and ending at *x*=1 and hitting *x*=0 at least once
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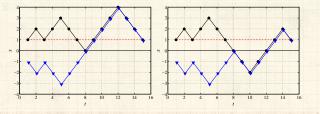
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of *t*-step paths starting and ending at *x*=1 and hitting *x*=0 at least once
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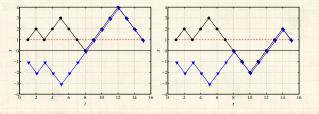
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Key observation for excluded walks:

- So For any path starting at x=1 that hits 0, there is a unique matching path starting at x=-1.
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So $N_{\text{first return}}(2n) = N(1, 1, 2n - 2) - N(-1, 1, 2n - 2)$

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Probability of first return: Insert question from assignment 5 📿 :

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Probability of first return: Insert question from assignment 5 🐼 : 🍪 Find

$$N_{\rm fr}(2n) \sim \frac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}.$$

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Normalized number of paths gives probability.
 Total number of possible paths = 2^{2n} .

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Probability of first return: Insert question from assignment 5 🗗 : 🇞 Find

2

$$N_{\rm fr}(2n) \sim \frac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}.$$

Normalized number of paths gives probability.
 Total number of possible paths = 2^{2n} .

$$P_{\mathsf{fr}}(2n) = \frac{1}{2^{2n}} N_{\mathsf{fr}}(2n)$$

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$$\simeq \frac{1}{2^{2n}} \frac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}$$

$$=\frac{1}{\sqrt{2\pi}}(2n)^{-3/2}$$

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Probability of first return: Insert question from assignment 5 🗗 : 🇞 Find

2

$$N_{\rm fr}(2n) \sim \frac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}.$$

Normalized number of paths gives probability.
 Total number of possible paths = 2^{2n} .

$$P_{\mathsf{fr}}(2n) = \frac{1}{2^{2n}} N_{\mathsf{fr}}(2n)$$

$$\simeq \frac{1}{2^{2n}} \frac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}$$

$$=\frac{1}{\sqrt{2\pi}}(2n)^{-3/2}\propto t^{-3/2}.$$

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion





 \clubsuit We have $P(t) \propto t^{-3/2}, \ \gamma = 3/2.$

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$$\red{We}$$
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6

line walks.

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$$\ref{eq:point_started}$$
 We have $P(t) \propto t^{-3/2}, \ \gamma = 3/2.$

Same scaling holds for continuous space/time walks. P(t) is normalizable. The PoCSverse Power-Law Mechanisms, Pt. 1 29 of 48

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$$\red{black}$$
 We have $P(t) \propto t^{-3/2}, \; \gamma = 3/2.$

- Same scaling holds for continuous space/time walks.
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- 🚳 Recurrence: Random walker always returns to origin



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- But mean, variance, and all higher moments are infinite.
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Higher dimensions 🕼:

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Higher dimensions 🕼:

- \Im Walker in d = 2 dimensions must also return
- $\,$ $\,$ Walker may not return in $d\geq 3$ dimensions
- 🚳 Associated human genius: George Pólya 🗹

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On finite spaces:

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On finite spaces:

In any finite homogeneous space, a random walker will visit every site with equal probability

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On finite spaces:

- In any finite homogeneous space, a random walker will visit every site with equal probability
- Call this probability the Invariant Density of a dynamical system

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On finite spaces:

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On networks:

- Equal probability still present: walkers traverse edges with equal frequency.

#totallygroovy

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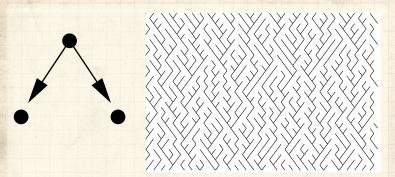
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Scheidegger Networks^[17, 4]



Random directed network on triangular lattice.
 Toy model of real networks.

'Flow' is southeast or southwest with equal probability.

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lacktrian creates basins with random walk boundaries.

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Creates basins with random walk boundaries.
 Observe that subtracting one random walk from another gives random walk with increments:

 $\epsilon_t = \left\{ \begin{array}{ll} +1 & \text{with probability } 1/4 \\ 0 & \text{with probability } 1/2 \\ -1 & \text{with probability } 1/4 \end{array} \right.$

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🚓 Random walk with probabilistic pauses.

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 Basin termination = first return random walk problem.

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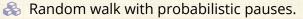
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Basin termination = first return random walk problem.

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🗞 Random walk with probabilistic pauses.

- Basin termination = first return random walk problem.
- \clubsuit Basin length ℓ distribution: $P(\ell) \propto \ell^{-3/2}$
- \clubsuit For real river networks, generalize to $P(\ell) \propto \ell^{-\gamma}$.

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rightarrow For a basin of length ℓ , width $\propto \ell^{1/2}$

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So For a basin of length ℓ , width $\propto \ell^{1/2}$ Basin area $a \propto \ell \cdot \ell^{1/2} = \ell^{3/2}$ The PoCSverse Power-Law Mechanisms, Pt. 1 33 of 48

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Solution For a basin of length ℓ , width $\propto \ell^{1/2}$ Basin area $a \propto \ell \cdot \ell^{1/2} = \ell^{3/2}$ Invert: $\ell \propto a^{2/3}$ The PoCSverse Power-Law Mechanisms, Pt. 1 33 of 48

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Both basin area and length obey power law distributions The PoCSverse Power-Law Mechanisms, Pt. 1 34 of 48

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- Both basin area and length obey power law distributions
- Observed for real river networks

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Generalize relationship between area and length:

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$$\ell \propto a^h$$
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So For real, large networks ^[13] $h \simeq 0.5$ (isometric scaling)

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Smaller basins possibly h > 1/2 (allometric scaling).

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- Solution Plan: Redo calc with γ , τ , and h.

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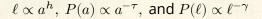
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$$\ell \propto a^h, \ P(a) \propto a^{-\tau}, \ \text{and} \ P(\ell) \propto \ell^{-\gamma}$$

 ${\clubsuit} \, \operatorname{d}\!\ell \propto \operatorname{d}\!(a^h) = ha^{h-1}\operatorname{d}\!a$

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2

$$\ell \propto a^h, \ P(a) \propto a^{-\tau}, \ \text{and} \ P(\ell) \propto \ell^{-\gamma}$$

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$$\tau = 1 + h(\gamma - 1)$$

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2

Given

$$\ell \propto a^h, \ P(a) \propto a^{-\tau}, \ \text{and} \ P(\ell) \propto \ell^{-\gamma}$$

$$\begin{aligned} & \& d\ell \propto d(a^h) = ha^{h-1} da \\ & \& \text{ Find } \tau \text{ in terms of } \gamma \text{ and } h. \\ & \& \text{ Pr}(\text{basin area} = a) da \\ & = \text{Pr}(\text{basin length} = \ell) d\ell \\ & \propto \ell^{-\gamma} d\ell \\ & \propto (a^h)^{-\gamma} a^{h-1} da \\ & = a^{-(1+h (\gamma-1))} da \end{aligned}$$

$$\tau = 1 + h(\gamma - 1)$$

Excellent example of the Scaling Relations found between exponents describing power laws for many systems.

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With more detailed description of network structure, $\tau = 1 + h(\gamma - 1)$ simplifies to: ^[3]

$$\tau = 2 - h$$

and



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- \mathfrak{S} Only one exponent is independent (take h).
- 🚳 Simplifies system description.
- Expect Scaling Relations where power laws are found.
- Need only characterize Universality C class with independent exponents.

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Death ...

Failure:

A very simple model of failure/death x_t = entity's 'health' at time t
Start with $x_0 > 0$.
Entity fails when x hits 0.

	And in case of the second seco

"Explaining mortality rate plateaus" Weitz and Fraser, Proc. Natl. Acad. Sci., **98**, 15383–15386, 2001. ^[18] The PoCSverse Power-Law Mechanisms, Pt. 1 37 of 48

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... and the NBA:

Basketball and other sports^[2]:



Three arcsine laws C (Lévy ^[12]) for continuous-time random walk last time T:

 $\overline{\pi} \sqrt{t(T-t)}$.

The arcsine distribution C applies for: (1) fraction of time positive, (2) the last time the walk changes sign, and (3) the time the maximum is achieved. \aleph Well approximated by basketball score lines ^[8, 2]. Australian Rules Football has some differences^[11]. The PoCSverse Power-Law Mechanisms, Pt. 1 38 of 48

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Can generalize to Fractional Random Walks^[15, 16, 14]

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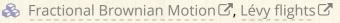
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 See Montroll and Shlesinger for example: ^[14]
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rightarrow In 1-d, standard deviation σ scales as

 $\sigma \sim t^{\alpha}$

 $\alpha = 1/2$ — diffusive $\alpha > 1/2$ — superdiffusive $\alpha < 1/2$ — subdiffusive

Extensive memory of path now matters...

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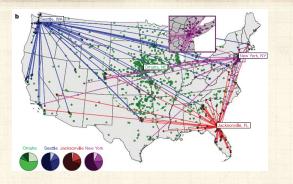
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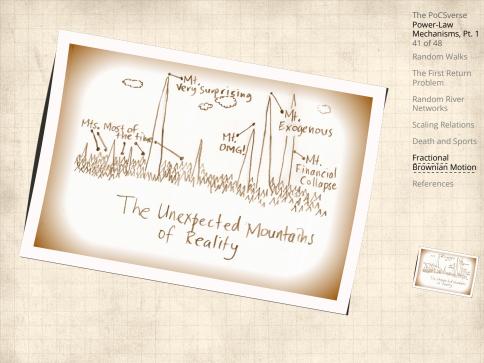
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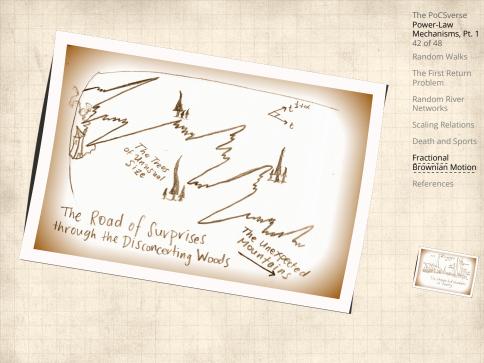
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- First big studies of movement and interactions of people.
- Brockmann *et al.* ^[1] "Where's George" study.
- Beyond Lévy: Superdiffusive in space but with long waiting times.
- Tracking movement via cell phones ^[9] and Twitter ^[7].





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