Mechanisms for Generating Power-Law Size Distributions, Part 1

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Principles of Complex Systems, Vols. 1 & 2 CSYS/MATH 300 and 303, 2021–2022 |@pocsvox

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The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion



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Outline

Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

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Random Walks

The First Return Problem

Random River Networks

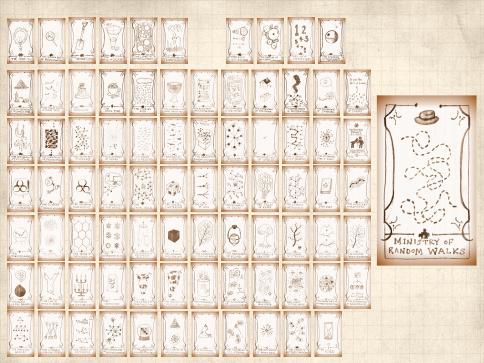
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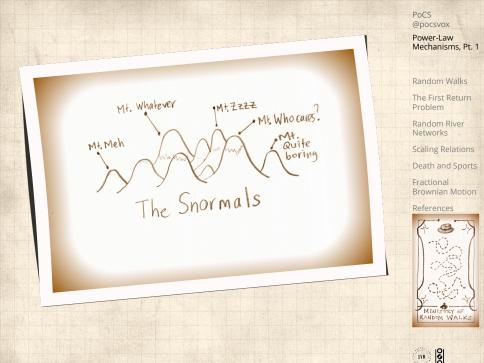
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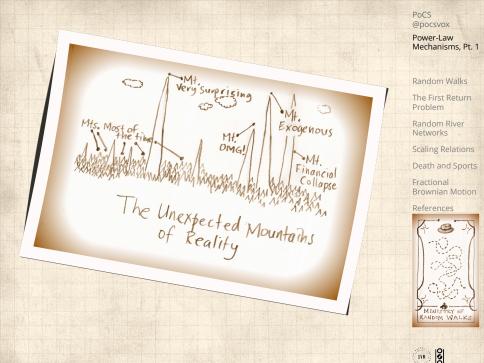
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Mechanisms:

A powerful story in the rise of complexity:
Structure arises out of randomness.
Structure A: Random walks.

The essential random walk:

- 🚳 One spatial dimension.
- 🚳 Time and space are discrete
- Random walker (e.g., a zombie texter \mathbb{C}) starts at origin x = 0.
- \clubsuit Step at time t is ϵ_t :

 $\epsilon_t = \left\{ \begin{array}{l} +1 & \text{with probability 1/2} \\ -1 & \text{with probability 1/2} \end{array} \right.$

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

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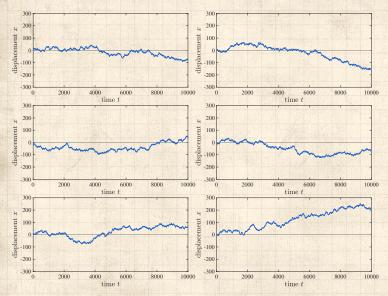
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A few random random walks:



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The First Return Problem

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Random walks:

6

Displacement after *t* steps:

$$x_t = \sum_{i=1}^t \epsilon_i$$

Expected displacement:

$$\langle x_t \rangle = \left\langle \sum_{i=1}^t \epsilon_i \right\rangle = \sum_{i=1}^t \left\langle \epsilon_i \right\rangle = 0$$

- Obviously fails for odd number of steps...
- But as time goes on, the chance of our texting undead friend lurching back to x=0 must diminish, right?

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The First Return Problem

Random River Networks

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Variances sum: 🗗

$$\mathsf{Var}(x_t) = \mathsf{Var}\left(\sum_{i=1}^t \epsilon_i\right)$$

$$=\sum_{i=1}^{\iota}\operatorname{Var}\left(\epsilon_{i}\right)=\sum_{i=1}^{\iota}1=t$$

* Sum rule = a good reason for using the variance to measure spread; only works for independent distributions.

So typical displacement from the origin scales as:

$$\sigma = t^{1/2}$$

A non-trivial scaling law arises out of additive aggregation or accumulation. PoCS @pocsvox

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Random River Networks

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Great moments in Televised Random Walks:

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The First Return Problem

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🗞 Also known as the bean machine 🗹, the quincunx (simulation), and the Galton box.

http://www.youtube.com/watch?v=05gqx6eSyO0?rel=0 Plinko! C from the Price is Right.





Random walk basics:

Counting random walks:

- Each specific random walk of length t appears with a chance $1/2^t$.
- We'll be more interested in how many random walks end up at the same place.
- Solution Define N(i, j, t) as # distinct walks that start at x = i and end at x = j after t time steps.
- Random walk must displace by +(j-i) after t steps.

🗞 Insert question from assignment 5 🗹

$$N(i,j,t) = \binom{t}{(t+j-i)/2}$$

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How does $P(x_t)$ behave for large t?

- $rac{1}{3}$ Take time t = 2n to help ourselves.
- $x_{2n} \in \{0, \pm 2, \pm 4, \dots, \pm 2n\}$
- x_{2n} is even so set $x_{2n} = 2k$.
- \bigotimes Using our expression N(i, j, t) with i = 0, j = 2k, and t = 2n, we have

$$\Pr(x_{2n} \equiv 2k) \propto \binom{2n}{n+k}$$

 \mathfrak{R} For large n, the binomial deliciously approaches the Normal Distribution of Snoredom:

$$\mathbf{Pr}(x_t \equiv x) \simeq \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}}$$

Insert question from assignment 5 🕑 The whole is different from the parts. #nutritious 😤 See also: Stable Distributions 🗹

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The First Return Problem

Random River Networks

Scaling Relations

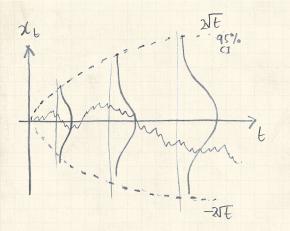
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Universality C is also not left-handed:



This is Diffusion C: the most essential kind of spreading (more later).

🚳 View as Random Additive Growth Mechanism.

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The First Return Problem

Random River Networks

Scaling Relations

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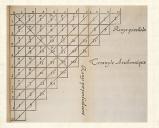
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So many things are connected:

Pascal's Triangle





Could have been the Pyramid of Pingala ³¹ or the Triangle of Khayyam, Jia Xian, Tartaglia, ...

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The First Return Problem

Random River Networks

Scaling Relations

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References



😣 Binomials tend towards the Normal.

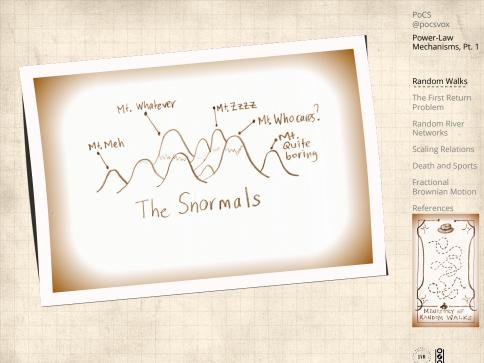
Counting encoded in algebraic forms (and much more).

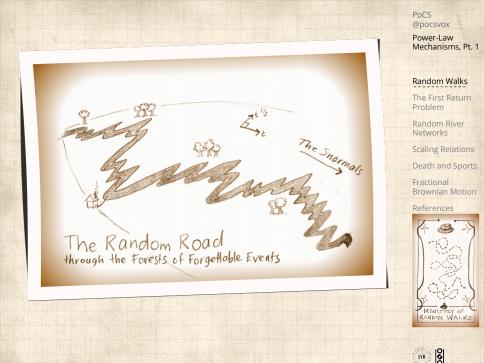
$$(h+t)^n = \sum_{k=0}^n \binom{n}{k} h^k t^{n-k} \text{ where } \binom{n}{k} = \frac{n!}{k!(n-k)!}$$

 $\label{eq:heat} \textcircled{\begin{subarray}{c} (h+t)^3 = hhh + hht + hth + thh + htt + tht + tth + ttt \\ \end{subarray}}$

¹Stigler's Law of Eponymy **C** showing excellent form again.







Random walks are even weirder than you might think...

- $\underset{r,t}{\Leftrightarrow} \xi_{r,t}$ = the probability that by time step t, a random walk has crossed the origin r times.
- A Think of a coin flip game with ten thousand tosses.
- If you are behind early on, what are the chances you will make a comeback?
- A The most likely number of lead changes is... 0.

$$\diamondsuit$$
 In fact: $\xi_{0,t} > \xi_{1,t} > \xi_{2,t} > \cdots$

\lambda Even crazier:

The expected time between tied scores = ∞

See Feller, Intro to Probability Theory, Volume I^[5]

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

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Fractional Brownian Motion





Applied knot theory:



"Designing tie knots by random walks" Fink and Mao, Nature, **398**, 31–32, 1999.^[6]





Figure 1 All diagrams are drawn in the frame of reference of the mirror image of the actual tie. **a**, The two ways of beginning a knot, $L_{\rm o}$ and $L_{\rm o}$. For knots beginning with $L_{\rm o}$, the tie must begin inside-out. **b**, The four-in-hand, denoted by the sequence $L_{\rm o} R_{\rm o} L_{\rm o} C_{\rm o} T$. **c**, A knot may be represented by a persistent random walk on a triangular lattice. The example shown is the four-in-hand, indicated by the walk 1116.

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

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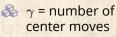
Applied knot theory:

Table 1 Aesthetic tie knots

h	γ	γ/h	K(h, γ)	S	b	Name	Sequence
3	1	0.33	1	0	0		L _☉ R _⊗ C _☉ T
4	1	0.25	1	- 1	1	Four-in-hand	$L_{\otimes}R_{\odot}L_{\otimes}C_{\odot}T$
5	2	0.40	2	- 1	0	Pratt knot	$L_{\odot}C_{\otimes}R_{\odot}L_{\otimes}C_{\odot}T$
6	2	0.33	4	0	0	Half-Windsor	$L_{\otimes}R_{\odot}C_{\otimes}L_{\odot}R_{\otimes}C_{\odot}T$
7	2	0.29	6	- 1	1		$L_{\odot}R_{\otimes}L_{\odot}C_{\otimes}R_{\odot}L_{\otimes}C_{\odot}T$
7	3	0.43	4	0	1		$L_{\odot}C_{\otimes}R_{\odot}C_{\otimes}L_{\odot}R_{\otimes}C_{\odot}T$
8	2	0.25	8	0	2		$L_{\otimes}R_{\odot}L_{\otimes}C_{\odot}R_{\otimes}L_{\odot}R_{\otimes}C_{\odot}T$
8	3	0.38	12	- 1	0	Windsor	$L_{\otimes}C_{\odot}R_{\otimes}L_{\odot}C_{\otimes}R_{\odot}L_{\otimes}C_{\odot}T$
9	3	0.33	24	0	0		$L_{\odot}R_{\otimes}C_{\odot}L_{\otimes}R_{\odot}C_{\otimes}L_{\odot}R_{\otimes}C_{\odot}T$
9	4	0.44	8	- 1	2		$L_{\circ}C_{\otimes}R_{\circ}C_{\otimes}L_{\circ}C_{\otimes}R_{\circ}L_{\otimes}C_{\circ}T$

Knots are characterized by half-winding number h, centre number γ , centre fraction γ/h , knots per class $K(h, \gamma)$, symmetry s, balance b, name and sequence.

h = number of moves



 $\stackrel{}{\bigotimes} \begin{array}{l} K(h,\gamma) = \\ 2^{\gamma-1} {h-\gamma-2 \choose \gamma-1} \end{array}$

 $s = \sum_{i=1}^{h} x_i \text{ where } x = -1$ for *L* and +1 for *R*.

 $\begin{array}{l} \textcircled{\&} \quad b = \frac{1}{2} \sum_{i=2}^{h-1} |\omega_i + \omega_{i-1}| \\ \text{where } \omega = \pm 1 \\ \text{represents winding} \\ \text{direction.} \end{array}$

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The First Return Problem

Random River Networks

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References

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Random walks #crazytownbananapants

The problem of first return:

- What is the probability that a random walker in one dimension returns to the origin for the first time after t steps?
- Will our zombie texter always return to the origin?
- What about higher dimensions?

Reasons for caring:

- 1. We will find a power-law size distribution with an interesting exponent.
- 2. Some physical structures may result from random walks.
- 3. We'll start to see how different scalings relate to each other.

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The First Return Problem

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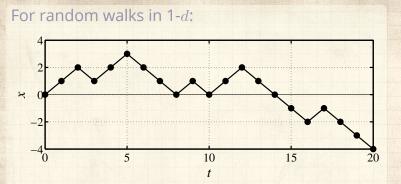
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- A return to origin can only happen when t = 2n. In example above, returns occur at t = 8, 10, and 14.
- \mathfrak{R} Call $P_{\mathsf{fr}(2n)}$ the probability of first return at t = 2n.
- Probability calculation \equiv Counting problem (combinatorics/statistical mechanics).
- Idea: Transform first return problem into an easier return problem.

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Random Walks

The First Return Problem

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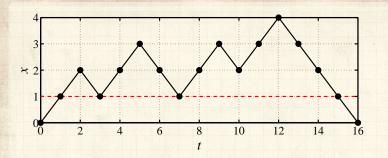
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Solution Can assume zombie texter first lurches to x = 1.

Solution Observe walk first returning at t = 16 stays at or above x = 1 for $1 \le t \le 15$ (dashed red line).

 \Im Now want walks that can return many times to x = 1.

 $\begin{array}{l} \textcircled{\&} \quad P_{\mathsf{fr}}(2n) = \\ 2 \cdot \frac{1}{2} Pr(x_t \geq 1, 1 \leq t \leq 2n-1, \text{ and } x_1 = x_{2n-1} = 1) \end{array} \\ \end{array}$

 $rac{1}{2}$ The $rac{1}{2}$ accounts for $x_{2n} = 2$ instead of 0.

 \Im The 2 accounts for texters that first lurch to x = -1.

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



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Counting first returns:

Approach:

- Move to counting numbers of walks.
- 🚳 Return to probability at end.
- Again, N(i, j, t) is the # of possible walks between x = i and x = j taking t steps.
- Solution Consider all paths starting at x = 1 and ending at x = 1 after t = 2n 2 steps.
- Solution Idea: If we can compute the number of walks that hit x = 0 at least once, then we can subtract this from the total number to find the ones that maintain $x \ge 1$.
- $rac{2}{3}$ Call walks that drop below x = 1 excluded walks.
- We'll use a method of images to identify these excluded walks.

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

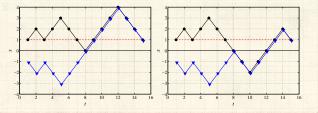
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Examples of excluded walks:



Key observation for excluded walks:

- So For any path starting at x=1 that hits 0, there is a unique matching path starting at x=-1.
- Matching path first mirrors and then tracks after first reaching x=0.

of t-step paths starting and ending at x=1 and hitting x=0 at least once = # of t-step paths starting at x=-1 and ending at x=1 = N(-1, 1, t)

So $N_{\text{first return}}(2n) = N(1, 1, 2n - 2) - N(-1, 1, 2n - 2)$

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



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Probability of first return: Insert question from assignment 5 C :

2

$$N_{\rm fr}(2n) \sim \frac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}.$$

Solution Normalized number of paths gives probability. Total number of possible paths = 2^{2n} .

$$P_{\mathsf{fr}}(2n) = \frac{1}{2^{2n}} N_{\mathsf{fr}}(2n)$$

$$\simeq \frac{1}{2^{2n}} \frac{2^{2n-3/2}}{\sqrt{2\pi}n^{3/2}}$$

$$=\frac{1}{\sqrt{2\pi}}(2n)^{-3/2}\propto t^{-3/2}.$$

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion





- $\ref{eq:point}$ We have $P(t) \propto t^{-3/2}, \ \gamma = 3/2.$
- Same scaling holds for continuous space/time walks.
- $\bigotimes P(t)$ is normalizable.
- 🚳 Recurrence: Random walker always returns to origin
- But mean, variance, and all higher moments are infinite.
 #totalmadness
- 🗞 Even though walker must return, expect a long wait...
- One moral: Repeated gambling against an infinitely wealthy opponent must lead to ruin.

Higher dimensions 🕼:

- \Im Walker in d = 2 dimensions must also return
- \circledast Walker may not return in $d \geq 3$ dimensions
- 🚳 Associated human genius: George Pólya 🗹

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



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Random walks

On finite spaces:

- In any finite homogeneous space, a random walker will visit every site with equal probability
- Call this probability the Invariant Density of a dynamical system
- Non-trivial Invariant Densities arise in chaotic systems.

On networks:

- Equal probability still present: walkers traverse edges with equal frequency. #totallygroovy

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

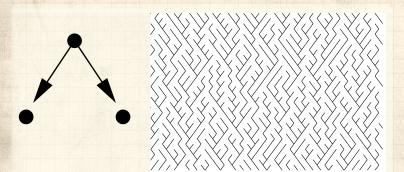
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Scheidegger Networks^[17, 4]



Random directed network on triangular lattice.
 Toy model of real networks.

'Flow' is southeast or southwest with equal probability.

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion





Scheidegger networks

Creates basins with random walk boundaries.
 Observe that subtracting one random walk from another gives random walk with increments:

 $\epsilon_t = \left\{ \begin{array}{ll} +1 & \text{with probability } 1/4 \\ 0 & \text{with probability } 1/2 \\ -1 & \text{with probability } 1/4 \end{array} \right.$

- 🗞 Random walk with probabilistic pauses.
- Basin termination = first return random walk problem.
- \clubsuit Basin length ℓ distribution: $P(\ell) \propto \ell^{-3/2}$
- \clubsuit For real river networks, generalize to $P(\ell) \propto \ell^{-\gamma}$.

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



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 $rac{2}{8}$ For a basin of length ℓ , width $\propto \ell^{1/2}$ Basin area $a \propto \ell \cdot \ell^{1/2} = \ell^{3/2}$ A Invert: $\ell \propto a^{2/3}$ $d\ell \propto d(a^{2/3}) = 2/3a^{-1/3}da$ \mathbf{R} **Pr**(basin area = a)da = **Pr**(basin length $= \ell$)d ℓ $\propto \ell^{-3/2} \mathsf{d} \ell$ $\propto (a^{2/3})^{-3/2}a^{-1/3}\mathsf{d}a$ $= a^{-4/3} da$ $= a^{-\tau} da$

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



na @ 33 of 48

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- Both basin area and length obey power law distributions
- 🚳 Observed for real river networks
- \clubsuit Reportedly: $1.3 < \tau < 1.5$ and $1.5 < \gamma < 2$

$$\ell \propto a^h$$

- So For real, large networks $^{[13]} h \simeq 0.5$ (isometric scaling)
- Smaller basins possibly h > 1/2 (allometric scaling).
- & Models exist with interesting values of h.
- Solution Plan: Redo calc with γ , τ , and h.

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



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2

Given

$$\ell \propto a^h, \ P(a) \propto a^{- au}, \ {
m and} \ P(\ell) \propto \ell^{-\gamma}$$

$$d\ell \propto d(a^{h}) = ha^{h-1}da$$

$$Find \tau \text{ in terms of } \gamma \text{ and } h.$$

$$Pr(\text{basin area} = a)da$$

$$= Pr(\text{basin length} = \ell)d\ell$$

$$\propto \ell^{-\gamma}d\ell$$

$$\propto (a^{h})^{-\gamma}a^{h-1}da$$

$$= a^{-(1+h(\gamma-1))}da$$

$$\tau = 1 + h(\gamma - 1)$$

Excellent example of the Scaling Relations found between exponents describing power laws for many systems. PoCS @pocsvox

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



990 35 of 48

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With more detailed description of network structure, $\tau = 1 + h(\gamma - 1)$ simplifies to: ^[3]

$$\tau = 2 - h$$

and

$$\gamma = 1/h$$

- \Im Only one exponent is independent (take h).
- 🗞 Simplifies system description.
- Expect Scaling Relations where power laws are found.
- Need only characterize Universality C class with independent exponents.

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



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Death ...

Failure:

A very simple model of failure/death x_t = entity's 'health' at time t
Start with $x_0 > 0$.
Entity fails when x hits 0.

"Explaining mortality rate plateaus" Weitz and Fraser, Proc. Natl. Acad. Sci., **98**, 15383–15386, 2001.^[18] PoCS @pocsvox

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

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References



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... and the NBA:

Basketball and other sports^[2]:

Three arcsine laws \mathbb{C} (Lévy ^[12]) for continuous-time random walk last time T:

 $\frac{1}{\pi} \frac{1}{\sqrt{t(T-t)}}.$

The arcsine distribution applies for:
(1) fraction of time positive, (2) the last time the walk changes sign, and (3) the time the maximum is achieved.
Well approximated by basketball score lines ^[8, 2].

Australian Rules Football has some differences [11].

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The First Return Problem

Random River Networks

Scaling Relations

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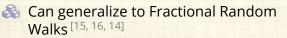
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More than randomness



Fractional Brownian Motion C, Lévy flights C
 See Montroll and Shlesinger for example: ^[14]
 "On 1/f noise and other distributions with long tails."

Proc. Natl. Acad. Sci., 1982.

rightarrow In 1-d, standard deviation σ scales as

 $\sigma \sim t^{\alpha}$

 $\alpha = 1/2$ — diffusive $\alpha > 1/2$ — superdiffusive $\alpha < 1/2$ — subdiffusive

🗞 Extensive memory of path now matters...

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

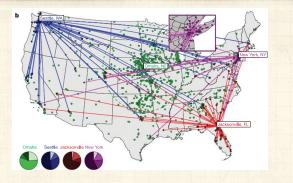
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- First big studies of movement and interactions of people.
- Brockmann *et al.* ^[1] "Where's George" study.
- Beyond Lévy: Superdiffusive in space but with long waiting times.

Tracking movement via cell phones^[9] and Twitter^[7]. PoCS @pocsvox

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

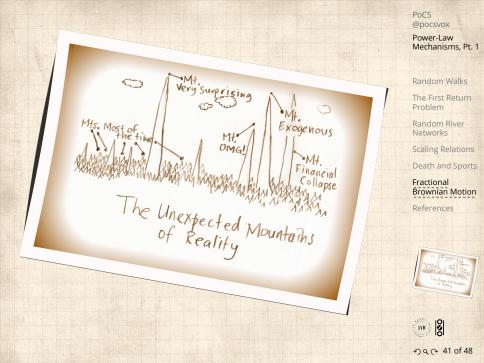
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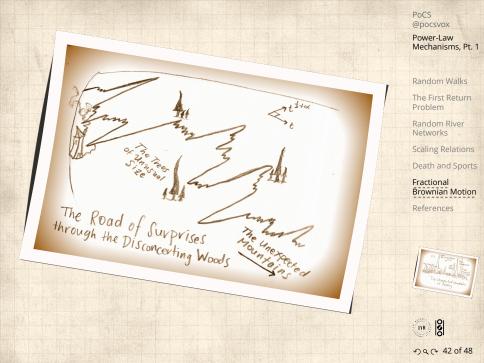
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Power-Law Mechanisms, Pt. 1

Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

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Power-Law Mechanisms, Pt. 1

Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

References



الله الح مرد 45 of 48

References IV

PoCS @pocsvox

Power-Law Mechanisms, Pt. 1

Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

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29 C 46 of 48

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The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

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Random Walks

The First Return Problem

Random River Networks

Scaling Relations

Death and Sports

Fractional Brownian Motion

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