

Mechanisms for Generating Power-Law Size Distributions, Part 1

Principles of Complex Systems | @pocsvox
CSYS/MATH 300, Fall, 2016 | #FallPoCS2016

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Vermont Advanced Computing Core | University of Vermont



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Power-Law
Mechanisms, Pt. 1

Sealie & Lambie
Productions



Random Walks
The First Return Problem
Examples

Variable
transformation
Basics
Holtmark's Distribution
PLIPLO

References



Outline

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Power-Law
Mechanisms, Pt. 1

Random Walks

The First Return Problem
Examples

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtsmark's Distribution

PLIPLO

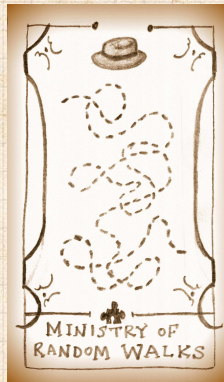
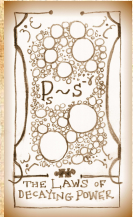
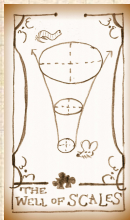
Variable transformation

Basics
Holtsmark's Distribution
PLIPLO

References

References





Random Walks

The First Return Problem
Examples

Variable transformation

Basics
Holtmark's Distribution
PLIPL0

References



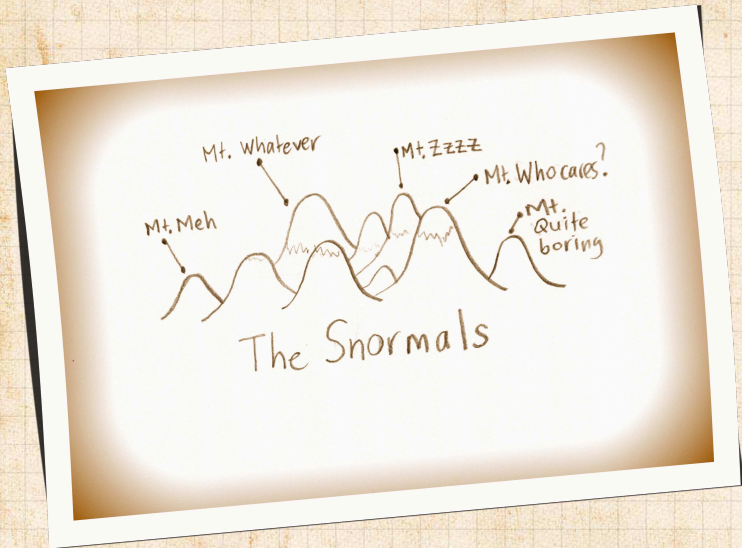
Random Walks

The First Return Problem
Examples

Variable
transformation

Basics
Holtmark's Distribution
PLIPL0

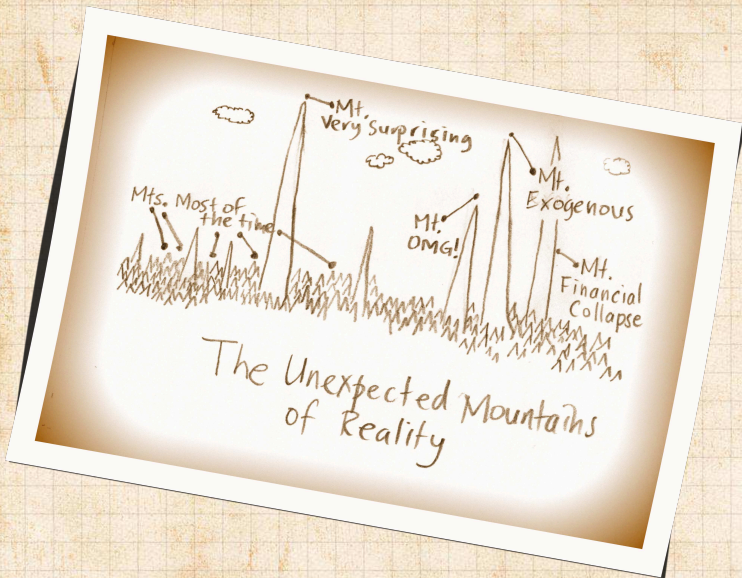
References



Random Walks
The First Return Problem
Examples

Variable
transformation
Basics
Holtzmark's Distribution
PLIPLO

References



Mechanisms:


A powerful story in the rise of complexity:


 structure arises out of randomness.


 Exhibit A: Random walks. 

The essential random walk:

 One spatial dimension.

 Time and space are discrete

 Random walker (e.g., a drunk) starts at origin
 $x = 0$.

 Step at time t is ϵ_t :

$$\epsilon_t = \begin{cases} +1 & \text{with probability } 1/2 \\ -1 & \text{with probability } 1/2 \end{cases}$$

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtmark's Distribution

PLIPLO

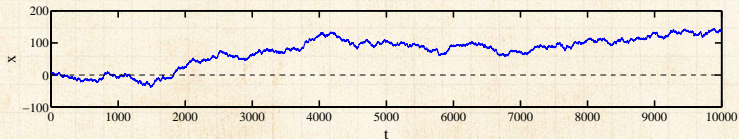
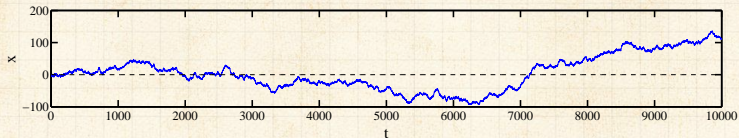
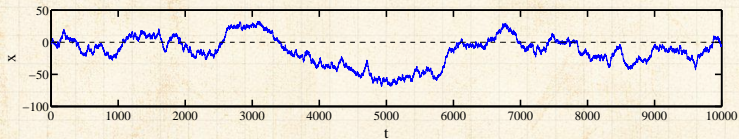
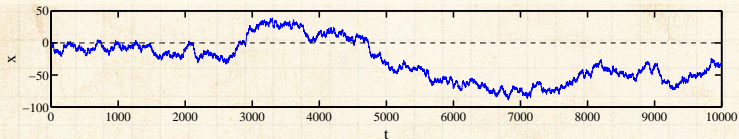
References



A few random random walks:

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Power-Law
Mechanisms, Pt. 1



Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtmark's Distribution

PLIPLO

References



Random walks:

Displacement after t steps:

$$x_t = \sum_{i=1}^t \epsilon_i$$

Expected displacement:

$$\langle x_t \rangle = \left\langle \sum_{i=1}^t \epsilon_i \right\rangle = \sum_{i=1}^t \langle \epsilon_i \rangle = 0$$

- At any time step, we 'expect' our drunkard to be back at the pub.
- Obviously fails for odd number of steps...
- But as time goes on, the chance of our drunkard lurching back to the pub must diminish, right?

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtmark's Distribution

PLIPLO

References



Variances sum: *

$$\begin{aligned}\text{Var}(x_t) &= \text{Var}\left(\sum_{i=1}^t \epsilon_i\right) \\ &= \sum_{i=1}^t \text{Var}(\epsilon_i) = \sum_{i=1}^t 1 = t\end{aligned}$$

* Sum rule = a good reason for using the variance to measure spread; only works for independent distributions.

So typical displacement from the origin scales as:

$$\sigma = t^{1/2}$$



A non-trivial scaling law arises out of additive aggregation or accumulation.

Random Walks

The First Return Problem

Examples

Variable

transformation

Basics

Holtzmark's Distribution

PLIPLO

References



Stock Market randomness:

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Power-Law
Mechanisms, Pt. 1

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtmark's Distribution

PLIPL0

References

Also known as the bean machine ↗, the quincunx
(simulation) ↗, and the Galton box.



Great moments in Televised Random Walks:

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtmark's Distribution

PLIPLD


References

Plinko! ↗ from the Price is Right.



Random walk basics:

Counting random walks:

- Each **specific** random walk of length t appears with a chance $1/2^t$.
- We'll be more interested in how many random walks end up at the same place.
- Define $N(i, j, t)$ as # distinct walks that start at $x = i$ and end at $x = j$ after t time steps.
- Random walk must displace by $+(j - i)$ after t steps.
- Insert question from assignment 3 

$$N(i, j, t) = \binom{t}{(t + j - i)/2}$$

Random Walks

The First Return Problem

Examples

Variable

transformation

Basics

Holtzmark's Distribution

PLIPLD

References



How does $P(x_t)$ behave for large t ?

Take time $t = 2n$ to help ourselves.

$x_{2n} \in \{0, \pm 2, \pm 4, \dots, \pm 2n\}$


x_{2n} is even so set $x_{2n} = 2k$.

Using our expression $N(i, j, t)$ with $i = 0, j = 2k$, and $t = 2n$, we have

$$\Pr(x_{2n} \equiv 2k) \propto \binom{2n}{n+k}$$

For large n , the binomial deliciously approaches the Normal Distribution of Snoredom:

$$\Pr(x_t \equiv x) \simeq \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}}$$

Insert question from assignment 3 

The whole is different from the parts. **#nutritious**

See also: Stable Distributions 

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtmark's Distribution

PLIPLD

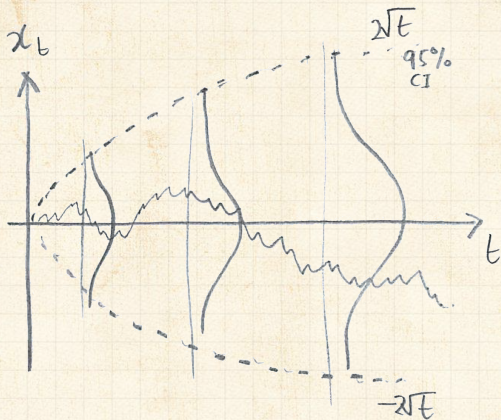
References



Universality is also not left-handed:

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Power-Law
Mechanisms, Pt. 1



Random Walks

The First Return Problem

Examples

Variable transformation


Basics

Holtmark's Distribution

PLIPL0

References



This is Diffusion : the most essential kind of spreading (more later).



View as Random Additive Growth Mechanism.



Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtsmark's Distribution

PLIPLD

References



The Random Road
through the Forests of Forgettable Events



Random Walks

The First Return Problem

Examples

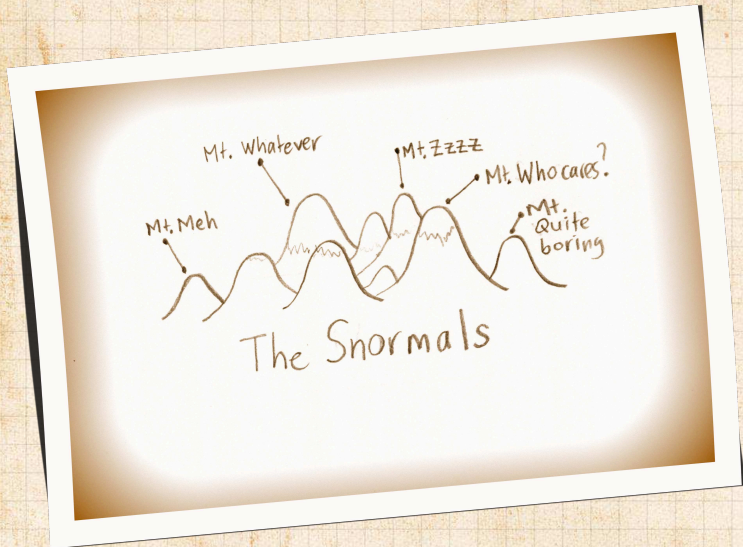
Variable transformation

Basics

Holtmark's Distribution

PLIPL0

References



Random walks are even weirder than you might think...

- 🧱 $\xi_{r,t}$ = the probability that by time step t , a random walk has crossed the origin r times.
- 🧱 Think of a coin flip game with ten thousand tosses.
- 🧱 If you are behind early on, what are the chances you will make a comeback?
- 🧱 The most likely number of lead changes is... 0.
- 🧱 In fact: $\xi_{0,t} > \xi_{1,t} > \xi_{2,t} > \dots$
- 🧱 Even crazier:
 The expected time between tied scores = ∞

See Feller, Intro to Probability Theory, Volume I [3]

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtzmark's Distribution


PLIPLD

References



Applied knot theory:



"Designing tie knots by random walks" 

Fink and Mao,
Nature, **398**, 31–32, 1999. [4]

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtzmark's Distribution

PLIPL0

References

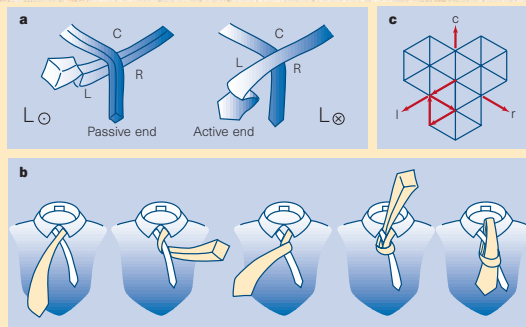


Figure 1 All diagrams are drawn in the frame of reference of the mirror image of the actual tie.


a. The two ways of beginning a knot, L_{\ominus} and L_{\otimes} . For knots beginning with L_{\ominus} , the tie must begin inside-out. **b.** The four-in-hand, denoted by the sequence $L_{\otimes}, R_{\ominus}, L_{\ominus}, C, T$. **c.** A knot may be represented by a persistent random walk on a triangular lattice. The example shown is the four-in-hand, indicated by the walk $\uparrow\uparrow\uparrow\uparrow$.





Table 1 **Aesthetic tie knots**


h	γ	γ/h	$K(h, \gamma)$	s	b	Name	Sequence
3	1	0.33	1	0	0		$L_0 R_0 C_0 T$
4	1	0.25	1	-1	1	Four-in-hand	$L_0 R_0 L_0 C_0 T$
5	2	0.40	2	-1	0	Pratt knot	$L_0 C_0 R_0 L_0 C_0 T$
6	2	0.33	4	0	0	Half-Windsor	$L_0 R_0 C_0 L_0 R_0 C_0 T$
7	2	0.29	6	-1	1		$L_0 R_0 L_0 C_0 R_0 L_0 C_0 T$
7	3	0.43	4	0	1		$L_0 C_0 R_0 C_0 L_0 R_0 C_0 T$
8	2	0.25	8	0	2		$L_0 R_0 L_0 C_0 R_0 L_0 R_0 C_0 T$
8	3	0.38	12	-1	0	Windsor	$L_0 C_0 R_0 L_0 C_0 R_0 L_0 C_0 T$
9	3	0.33	24	0	0		$L_0 R_0 C_0 L_0 R_0 C_0 L_0 R_0 C_0 T$
9	4	0.44	8	-1	2		$L_0 C_0 R_0 C_0 L_0 C_0 R_0 L_0 C_0 T$


Knots are characterized by half-winding number h , centre number γ , centre fraction γ/h , knots per class $K(h, \gamma)$, symmetry s , balance b , name and sequence.

 h = number of moves

 γ = number of center moves

 $K(h, \gamma) = 2^{\gamma-1} \binom{h-\gamma-2}{\gamma-1}$

 $s = \sum_{i=1}^h x_i$ where $x = -1$ for L and $+1$ for R .

 $b = \frac{1}{2} \sum_{i=2}^{h-1} |\omega_i + \omega_{i-1}|$ where $\omega = \pm 1$ represents winding direction.

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtmark's Distribution

PLIPLO

References



The problem of first return:

- What is the probability that a random walker in one dimension returns to the origin for the first time after t steps?
- Will our drunkard always return to the origin?
- What about higher dimensions?

Reasons for caring:

- We will find a power-law size distribution with an interesting exponent.
- Some physical structures may result from random walks.
- We'll start to see how different scalings relate to each other.

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

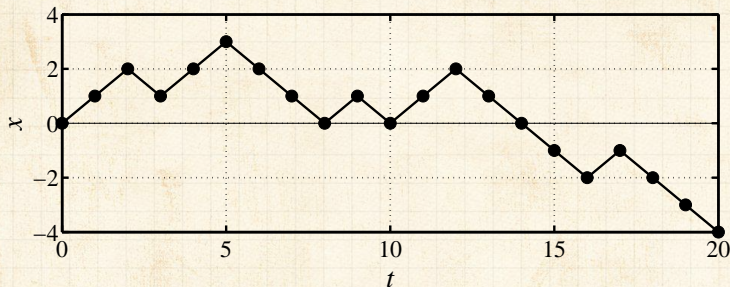
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References



For random walks in 1-d:



- 🧱 A **return** to origin can only happen when $t = 2n$.
- 🧱 In example above, returns occur at $t = 8, 10,$ and 14 .
- 🧱 Call $P_{fr(2n)}$ the probability of **first return** at $t = 2n$.
- 🧱 Probability calculation \equiv Counting problem (combinatorics/statistical mechanics).
- 🧱 **Idea:** Transform first return problem into an easier return problem.

Random Walks

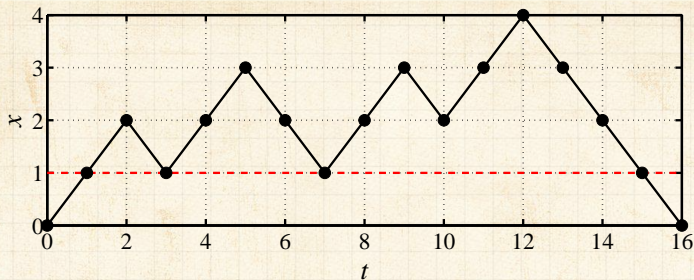
The First Return Problem
Examples

Variable transformation

Basics
Holtmark's Distribution
PLIPLD

References





- Can assume drunkard first lurches to $x = 1$.
- Observe walk first returning at $t = 16$ stays at or above $x = 1$ for $1 \leq t \leq 15$ (dashed red line).
- Now want walks that can return many times to $x = 1$.
- $P_{\text{fr}}(2n) = 2 \cdot \frac{1}{2} \text{Pr}(x_t \geq 1, 1 \leq t \leq 2n - 1, \text{ and } x_1 = x_{2n-1} = 1)$
- The $\frac{1}{2}$ accounts for $x_{2n} = 2$ instead of 0.
- The 2 accounts for drunkards that first lurch to $x = -1$.



Counting first returns:

Approach:

- Move to counting numbers of walks.
- Return to probability at end.
- Again, $N(i, j, t)$ is the # of possible walks between $x = i$ and $x = j$ taking t steps.
- Consider **all paths** starting at $x = 1$ and ending at $x = 1$ after $t = 2n - 2$ steps.
- Idea:** If we can compute the number of walks that hit $x = 0$ at least once, then we can subtract this from the total number to find the ones that maintain $x \geq 1$.
- Call walks that drop below $x = 1$ **excluded walks**.
- We'll use a method of images to identify these excluded walks.

Random Walks

The First Return Problem
Examples

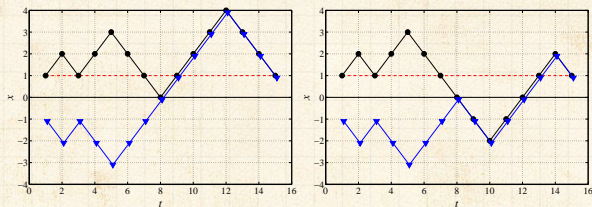
Variable transformation

Basics
Holtzmark's Distribution
PLIPLO

References



Examples of excluded walks:



Key observation for excluded walks:

- For any path starting at $x=1$ that hits 0, there is a unique matching path starting at $x=-1$.
- Matching path first mirrors and then tracks after first reaching $x=0$.
- # of t -step paths starting and ending at $x=1$ and hitting $x=0$ at least once
= # of t -step paths starting at $x=-1$ and ending at $x=1$ = $N(-1, 1, t)$
- So $N_{\text{first return}}(2n) = N(1, 1, 2n - 2) - N(-1, 1, 2n - 2)$





Probability of first return:

Insert question from assignment 3  :

 Find

$$N_{\text{fr}}(2n) \sim \frac{2^{2n-3/2}}{\sqrt{2\pi n^{3/2}}}.$$

 Normalized number of paths gives probability.

 Total number of possible paths = 2^{2n} .



$$\begin{aligned} P_{\text{fr}}(2n) &= \frac{1}{2^{2n}} N_{\text{fr}}(2n) \\ &\simeq \frac{1}{2^{2n}} \frac{2^{2n-3/2}}{\sqrt{2\pi n^{3/2}}} \\ &= \frac{1}{\sqrt{2\pi}} (2n)^{-3/2} \propto t^{-3/2}. \end{aligned}$$

Random Walks








The First Return Problem
Examples

Variable
transformation



Basics
Holtzmark's Distribution
PLIPL0

References



-  We have $P(t) \propto t^{-3/2}$, $\gamma = 3/2$.
-  Same scaling holds for continuous space/time walks.
-  $P(t)$ is normalizable.
-  **Recurrence:** Random walker always returns to origin
-  But mean, variance, and all higher moments are infinite. #totalmadness
-  Even though walker must return, expect a long wait...
-  **One moral:** Repeated gambling against an infinitely wealthy opponent must lead to ruin.

Higher dimensions

-  Walker in $d = 2$ dimensions must also return
-  Walker may not return in $d \geq 3$ dimensions

Random Walks

The First Return Problem
Examples

Variable transformation

Basics
Holtzmark's Distribution
PLIPLO

References



Random walks

On finite spaces:

- 🧱 In any finite homogeneous space, a random walker will visit every site with equal probability
- 🧱 Call this probability the **Invariant Density** of a dynamical system
- 🧱 Non-trivial Invariant Densities arise in chaotic systems.

On networks:

- 🧱 On networks, a random walker visits each node with frequency \propto node degree **#groovy**
- 🧱 Equal probability still present: walkers traverse **edges** with equal frequency. **#totallygroovy**

Random Walks

The First Return Problem

Examples

Variable transformation

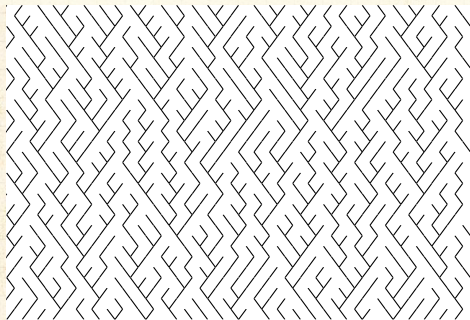
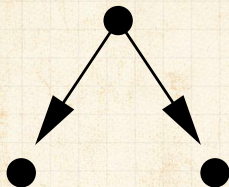
Basics




Holtmark's Distribution

PLIPLO

References





-  Random directed network on triangular lattice.
-  Toy model of real networks.
-  'Flow' is southeast or southwest with equal probability.

Random Walks

The First Return Problem

Examples

Variable transformation


Basics


Holtmark's Distribution

PLIPLD


References





 Creates basins with random walk boundaries.


 **Observe** that subtracting one random walk from another gives random walk with increments:

$$\epsilon_t = \begin{cases} +1 & \text{with probability } 1/4 \\ 0 & \text{with probability } 1/2 \\ -1 & \text{with probability } 1/4 \end{cases}$$

 Random walk with probabilistic pauses.

 Basin termination = first return random walk problem.

 Basin length ℓ distribution: $P(\ell) \propto \ell^{-3/2}$

 For real river networks, generalize to $P(\ell) \propto \ell^{-\gamma}$.

Random Walks

The First Return Problem

Examples

Variable

transformation

Basics


Holtmark's Distribution


PLIPLO


References





Connections between exponents:

 For a basin of length l , width $\propto l^{1/2}$

 Basin area $a \propto l \cdot l^{1/2} = l^{3/2}$

 Invert: $l \propto a^{2/3}$

 $dl \propto d(a^{2/3}) = 2/3 a^{-1/3} da$

 $\Pr(\text{basin area} = a) da$
 $= \Pr(\text{basin length} = l) dl$
 $\propto l^{-3/2} dl$
 $\propto (a^{2/3})^{-3/2} a^{-1/3} da$
 $= a^{-4/3} da$
 $= a^{-\tau} da$

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

Holtmark's Distribution

PLIPLO

References



Connections between exponents:

- Both basin area and length obey power law distributions
- Observed for real river networks
- Reportedly: $1.3 < \tau < 1.5$ and $1.5 < \gamma < 2$

Generalize relationship between area and length:

- Hack's law^[5]:

$$l \propto a^h.$$

- For real, large networks $h \simeq 0.5$
- Smaller basins possibly $h > 1/2$ (see: allometry).
- Models exist with interesting values of h .
- Plan: Redo calc with γ , τ , and h .

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtmark's Distribution

PLIPLD


References





Connections between exponents:

 Given

$$\ell \propto a^h, P(a) \propto a^{-\tau}, \text{ and } P(\ell) \propto \ell^{-\gamma}$$


 $d\ell \propto d(a^h) = ha^{h-1}da$

 Find τ in terms of γ and h .

 $\Pr(\text{basin area} = a)da$
 $= \Pr(\text{basin length} = \ell)d\ell$
 $\propto \ell^{-\gamma}d\ell$
 $\propto (a^h)^{-\gamma}a^{h-1}da$
 $= a^{-(1+h(\gamma-1))}da$



$$\tau = 1 + h(\gamma - 1)$$

 Excellent example of the **Scaling Relations** found between exponents describing power laws for many systems.

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtmark's Distribution

PLIPLO

References








Connections between exponents:

With more detailed description of network structure, $\tau = 1 + h(\gamma - 1)$ simplifies to: ^[1]

$$\tau = 2 - h$$

and

$$\gamma = 1/h$$

-  Only one exponent is independent (take h).
-  Simplifies system description.
-  Expect Scaling Relations where power laws are found.
-  Need only characterize Universality  class with independent exponents.

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtmark's Distribution


PLIPLD


References





Other First Returns or First Passage Times:

Failure:


 A very simple model of failure/death: ^[11]


 x_t = entity's 'health' at time t

 Start with $x_0 > 0$.

 Entity fails when x hits 0.

Streams

 Dispersion of suspended sediments in streams.

 Long times for clearing.



More than randomness

Can generalize to Fractional Random Walks [7, 8, 6]

Levy flights, Fractional Brownian Motion

See Montroll and Shlesinger for example: [6]
"On $1/f$ noise and other distributions with long tails."

Proc. Natl. Acad. Sci., 1982.

In 1-d, standard deviation σ scales as

$$\sigma \sim t^\alpha$$

$\alpha = 1/2$ — diffusive

$\alpha > 1/2$ — superdiffusive

$\alpha < 1/2$ — subdiffusive

Extensive memory of path now matters...

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtzmark's Distribution

PLIPL0

References



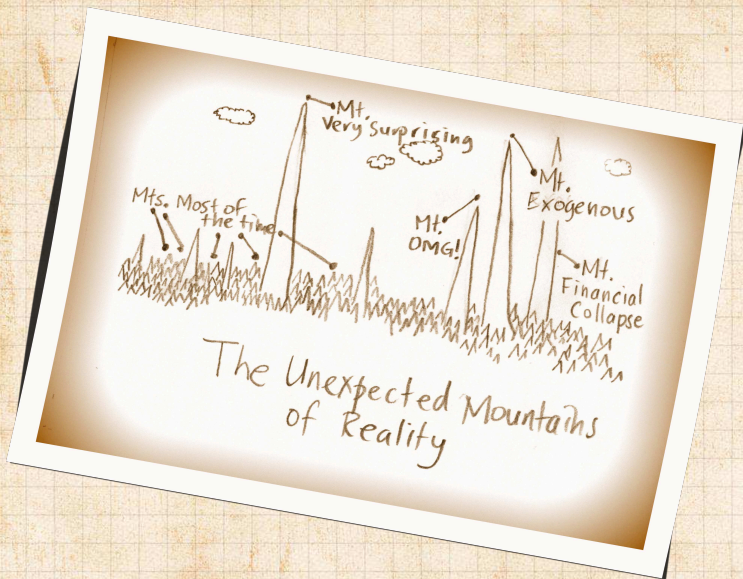
Random Walks
The First Return Problem

Examples

Variable
transformation

Basics
Holtzmark's Distribution
PLIPLO

References



Random Walks

The First Return Problem

Examples

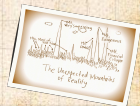
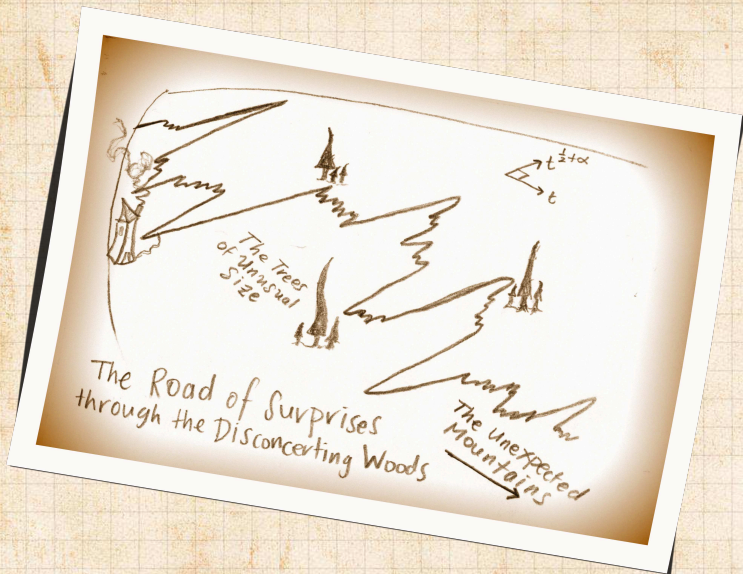
Variable
transformation

Basics

Holtzmark's Distribution

PLIPLD

References



Neural reboot (NR):

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Power-Law
Mechanisms, Pt. 1

Desert rain frog/Squeaky toy:

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

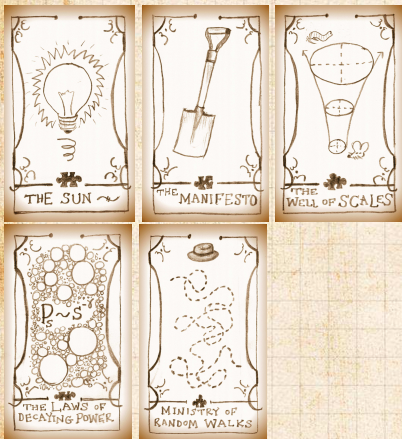
Holtmark's Distribution

PLIPLD

References

<https://www.youtube.com/v/cBkWhkAZ9ds?rel=0>





Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtmark's Distribution

PLIPL0

References



Understand power laws as arising from

1. Elementary distributions (e.g., exponentials).
2. Variables connected by power relationships.

🧩 Random variable X with known distribution P_x

🧩 Second random variable Y with $y = f(x)$.

$$\begin{aligned} \text{🧩 } P_Y(y)dy &= \\ &= \sum_{x|f(x)=y} P_X(x)dx \\ &= \sum_{y|f(x)=y} P_X(f^{-1}(y)) \frac{dy}{|f'(f^{-1}(y))|} \end{aligned}$$

🧩 Often easier to do by hand...

Random Walks

The First Return Problem
Examples

Variable
transformation

Basics
.....
Holtmark's Distribution
PLIPLO

References



General Example

Assume relationship between x and y is 1-1.

Power-law relationship between variables:

$$y = cx^{-\alpha}, \alpha > 0$$

Look at y large and x small



$$dy = d(cx^{-\alpha})$$

$$= c(-\alpha)x^{-\alpha-1}dx$$

$$\text{invert: } dx = \frac{-1}{c\alpha}x^{\alpha+1}dy$$

$$dx = \frac{-1}{c\alpha} \left(\frac{y}{c}\right)^{-(\alpha+1)/\alpha} dy$$

$$dx = \frac{-c^{1/\alpha}}{\alpha} y^{-1-1/\alpha} dy$$



Now make transformation:

$$P_y(y)dy = P_x(x)dx$$

$$P_y(y)dy = P_x \left(\overbrace{\left(\frac{y}{c}\right)^{-1/\alpha}}^{(x)} \right) \overbrace{\frac{c^{1/\alpha}}{\alpha} y^{-1-1/\alpha} dy}^{dx}$$

🧱 If $P_x(x) \rightarrow$ non-zero constant as $x \rightarrow 0$ then

$$P_y(y) \propto y^{-1-1/\alpha} \text{ as } y \rightarrow \infty.$$

🧱 If $P_x(x) \rightarrow x^\beta$ as $x \rightarrow 0$ then

$$P_y(y) \propto y^{-1-1/\alpha-\beta/\alpha} \text{ as } y \rightarrow \infty.$$





Example

Exponential distribution

Given $P_x(x) = \frac{1}{\lambda}e^{-x/\lambda}$ and $y = cx^{-\alpha}$, then

$$P(y) \propto y^{-1-1/\alpha} + O(y^{-1-2/\alpha})$$

 Exponentials arise from randomness (easy)...

 More later when we cover robustness.

Random Walks

The First Return Problem
Examples

Variable
transformation

Basics
.....
Holtzmark's Distribution
PLIPL0

References



Gravity

PoCS | @pocsvox

Power-Law
Mechanisms, Pt. 1



Select a random point in the universe \vec{x}



Measure the force of gravity $F(\vec{x})$



Observe that $P_F(F) \sim F^{-5/2}$.



Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtmark's Distribution

PLIPLO

References



Matter is concentrated in stars: ^[10]

☰ F is distributed unevenly

☰ Probability of being a distance r from a single star at $\vec{x} = \vec{0}$:

$$P_r(r)dr \propto r^2 dr$$

☰ Assume stars are distributed randomly in space (oops?)

☰ Assume only one star has significant effect at \vec{x} .

☰ Law of gravity:

$$F \propto r^{-2}$$

☰ invert:

$$r \propto F^{-1/2}$$

☰ Also invert:

$$dF \propto d(r^{-2}) \propto r^{-3} dr \rightarrow dr \propto r^3 dF \propto F^{-3/2} dF.$$

Random Walks

The First Return Problem
ExamplesVariable
transformationBasics
Holtmark's Distribution
PLIPLO

References



Transformation:

Using $r \propto F^{-1/2}$, $dr \propto F^{-3/2}dF$, and $P_r(r) \propto r^2$



$$P_F(F)dF = P_r(r)dr$$



$$\propto P_r(\text{const} \times F^{-1/2})F^{-3/2}dF$$



$$\propto (F^{-1/2})^2 F^{-3/2}dF$$



$$= F^{-1-3/2}dF$$



$$= F^{-5/2}dF.$$

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtmark's Distribution

PLIPLD

References



Gravity:

$$P_F(F) = F^{-5/2} dF$$

$$\gamma = 5/2$$



Mean is finite.



Variance = ∞ .



A wild distribution.



Upshot: Random sampling of space usually safe
but can end badly...

Random Walks

The First Return Problem

Examples

Variable

transformation

Basics

Holtmark's Distribution

PLIPLO

References



Doctorin' the Tardis

Random Walks

The First Return Problem

Examples

Variable transformation

Basics

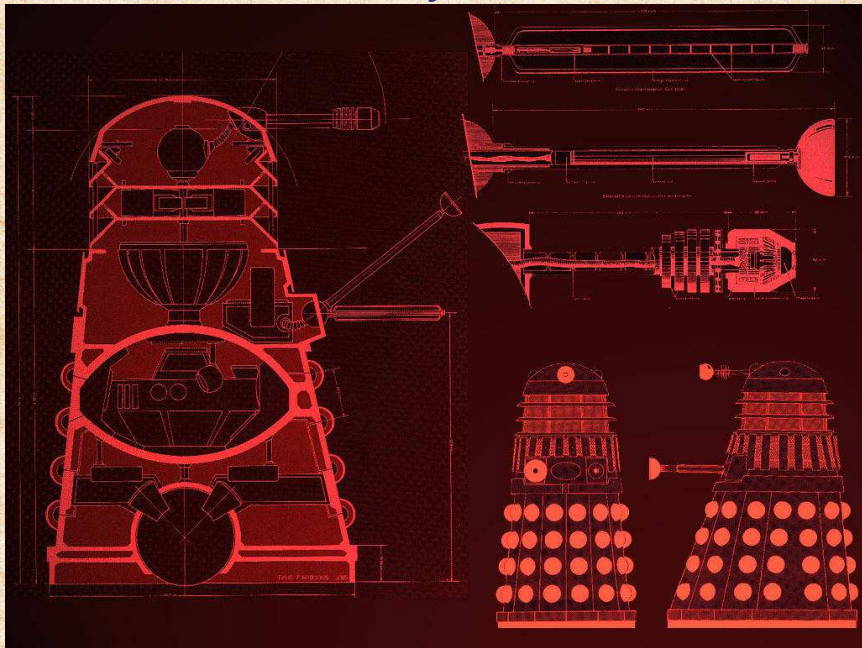
Holtmark's Distribution

PLIPL0

References



□ Todo: Build Dalek army.



Extreme Caution!

Random Walks

The First Return Problem
Examples

Variable transformation

Basics
Holtsmark's Distribution
PLIPLO

References

- PLIPLO = **Power law in, power law out**
- Explain a power law as resulting from another unexplained power law.
- Yet another homunculus argument...
- Don't do this!!! (slap, slap)
- MIWO = **Mild in, Wild out** is fine.
- In general: We need mechanisms!



Neural reboot (NR):

PoCS | @pocsvox

Power-Law
Mechanisms, Pt. 1

Zoomage in slow motion

Random Walks

The First Return Problem

Examples

Variable
transformation

Basics

Holtmark's Distribution

PLIPLO

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Random Walks
The First Return Problem
Examples

Variable
transformation
Basics
Holtzmark's Distribution
PLIPLO

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Random Walks

The First Return Problem
Examples

Variable
transformation

Basics
Holtzmark's Distribution
PLIPLO

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Random Walks
The First Return Problem
Examples

Variable
transformation
Basics
Holtmark's Distribution
PLIPL0

References

