

Lognormals and friends

Principles of Complex Systems | @pocsvox
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Prof. Peter Dodds | @peterdodds

Dept. of Mathematics & Statistics | Vermont Complex Systems Center
Vermont Advanced Computing Core | University of Vermont

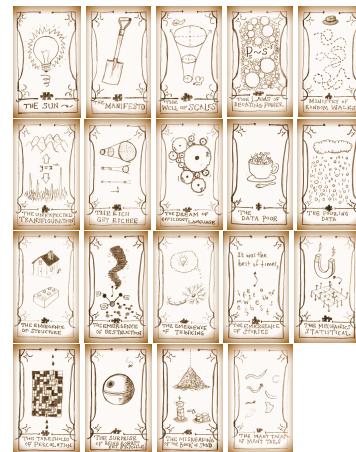


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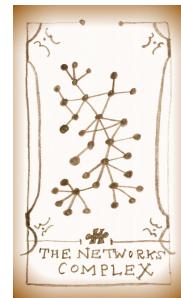


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Outline

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Alternative distributions

There are other ‘heavy-tailed’ distributions:

1. The Log-normal distribution ↗

$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right)$$

2. Weibull distributions ↗

$$P(x)dx = \frac{k}{\lambda} \left(\frac{x}{\lambda}\right)^{\mu-1} e^{-(x/\lambda)^{\mu}} dx$$

CCDF = stretched exponential ↗.

3. Gamma distributions ↗, and more.

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The lognormal distribution:

$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right)$$

⌚ In x is distributed according to a normal distribution with mean μ and variance σ .

⌚ Appears in economics and biology where growth increments are distributed normally.



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- Standard form reveals the mean μ and variance σ^2 of the underlying normal distribution:

$$P(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right)$$

- For lognormals:

$$\mu_{\text{lognormal}} = e^{\mu + \frac{1}{2}\sigma^2}, \quad \text{median}_{\text{lognormal}} = e^\mu,$$

$$\sigma_{\text{lognormal}} = (e^{\sigma^2} - 1)e^{2\mu + \sigma^2}, \quad \text{mode}_{\text{lognormal}} = e^{\mu - \sigma^2}.$$

- All moments of lognormals are finite.

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Confusion

What's happening:

$$\ln P(x) = \ln \left\{ \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right) \right\}$$

$$= -\ln x - \ln \sqrt{2\pi}\sigma - \frac{(\ln x - \mu)^2}{2\sigma^2}$$

$$= -\frac{1}{2\sigma^2} (\ln x)^2 + \left(\frac{\mu}{\sigma^2} - 1 \right) \ln x - \ln \sqrt{2\pi}\sigma - \frac{\mu^2}{2\sigma^2}.$$

If the first term is relatively small,

$$\ln P(x) \sim -\left(1 - \frac{\mu}{\sigma^2}\right) \ln x + \text{const.} \Rightarrow \gamma = 1 - \frac{\mu}{\sigma^2}$$



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Derivation from a normal distribution

Take Y as distributed normally:



$$P(y)dy = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(y - \mu)^2}{2\sigma^2}\right) dy$$

Set $Y = \ln X$:

- Transform according to $P(x)dx = P(y)dy$:



$$\frac{dy}{dx} = 1/x \Rightarrow dy = dx/x$$



$$\Rightarrow P(x)dx = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln x - \mu)^2}{2\sigma^2}\right) dx$$

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Confusion

- If $\mu < 0, \gamma > 1$ which is totally cool.

- If $\mu > 0, \gamma < 1$, not so much.

- If $\sigma^2 \gg 1$ and μ ,

$$\ln P(x) \sim -\ln x + \text{const.}$$



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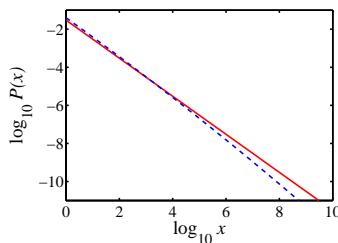
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Confusion between lognormals and pure power laws



Near agreement
over four orders
of magnitude!

- For lognormal (blue), $\mu = 0$ and $\sigma = 10$.
- For power law (red), $\gamma = 1$ and $c = 0.03$.

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Generating lognormals:



$$x_{n+1} = rx_n$$

where $r > 0$ is a random growth variable

- (Shrinkage is allowed)

- In log space, growth is by addition:

$$\ln x_{n+1} = \ln r + \ln x_n$$

- $\Rightarrow \ln x_n$ is normally distributed

- $\Rightarrow x_n$ is lognormally distributed



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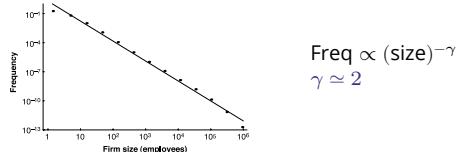


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Lognormals or power laws?

- Gibrat [2] (1931) uses preceding argument to explain lognormal distribution of firm sizes ($\gamma \approx 1$).
- But Robert Axtell [1] (2001) shows a power law fits the data very well with $\gamma = 2$, not $\gamma = 1$ (!)
- Problem of data censusing (missing small firms).



- One piece in Gibrat's model seems okay empirically: Growth rate r appears to be independent of firm size. [1].

An explanation

- Axtell cites Malcai et al.'s (1999) argument [5] for why power laws appear with exponent $\gamma \approx 2$
- The set up: N entities with size $x_i(t)$
- Generally:

$$x_i(t+1) = rx_i(t)$$

- where r is drawn from some happy distribution
- Same as for lognormal but one extra piece.
 - Each x_i cannot drop too low with respect to the other sizes:

$$x_i(t+1) = \max(rx_i(t), c \langle x_i \rangle)$$

Some math later...

Insert question from assignment 7 ↗

- Find $P(x) \sim x^{-\gamma}$

- where γ is implicitly given by

$$N = \frac{(\gamma - 2)}{(\gamma - 1)} \left[\frac{(c/N)^{\gamma-1} - 1}{(c/N)^{\gamma-1} - (c/N)} \right]$$

N = total number of firms.



$$\text{Now, if } c/N \ll 1 \text{ and } \gamma > 2 \quad N = \frac{(\gamma - 2)}{(\gamma - 1)} \left[\frac{-1}{-(c/N)} \right]$$



$$\text{Which gives } \gamma \sim 1 + \frac{1}{1-c}$$

- Groovy... c small $\Rightarrow \gamma \approx 2$

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The second tweak

Ages of firms/people/... may not be the same

- Allow the number of updates for each size x_i to vary
- Example: $P(t)dt = ae^{-at}dt$ where t = age.
- Back to no bottom limit: each x_i follows a lognormal
- Sizes are distributed as [6]

$$P(x) = \int_{t=0}^{\infty} ae^{-at} \frac{1}{x\sqrt{2\pi t}} \exp\left(-\frac{(\ln x - \mu)^2}{2t}\right) dt$$

(Assume for this example that $\sigma \sim t$ and $\mu = \ln m$)

- Now averaging different lognormal distributions.

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Averaging lognormals

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$$P(x) = \int_{t=0}^{\infty} ae^{-at} \frac{1}{x\sqrt{2\pi t}} \exp\left(-\frac{(\ln \frac{x}{m})^2}{2t}\right) dt$$

- Insert fabulous calculation (team is spared).
- Some enjoyable suffering leads to:

$$P(x) \propto x^{-1} e^{-\sqrt{2\lambda}(\ln \frac{x}{m})^2}$$

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$$P(x) \propto x^{-1} e^{-\sqrt{2\lambda}(\ln \frac{x}{m})^2}$$

- Depends on sign of $\ln \frac{x}{m}$, i.e., whether $\frac{x}{m} > 1$ or $\frac{x}{m} < 1$.

$$P(x) \propto \begin{cases} x^{-1+\sqrt{2\lambda}} & \text{if } \frac{x}{m} < 1 \\ x^{-1-\sqrt{2\lambda}} & \text{if } \frac{x}{m} > 1 \end{cases}$$

- 'Break' in scaling (not uncommon)

- Double-Pareto distribution ↗

- First noticed by Montroll and Shlesinger [7, 8]

- Later: Huberman and Adamic [3, 4]: Number of pages per website

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Summary of these exciting developments:

- Lognormals and power laws can be **awfully** similar
- Random Multiplicative Growth leads to lognormal distributions
- Enforcing a minimum size leads to a power law tail
- With no minimum size but a distribution of lifetimes, the double Pareto distribution appears
- Take-home message: Be careful out there...

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